

Introduction to Multilevel Modeling Using HLM 6

By ATS Statistical Consulting Group

Multilevel data structure

- Students nested within schools
- Children nested within families
- Respondents nested within interviewers
- Repeated measures nested within individuals – longitudinal data, growth curve modeling

In the example of student nested within schools:

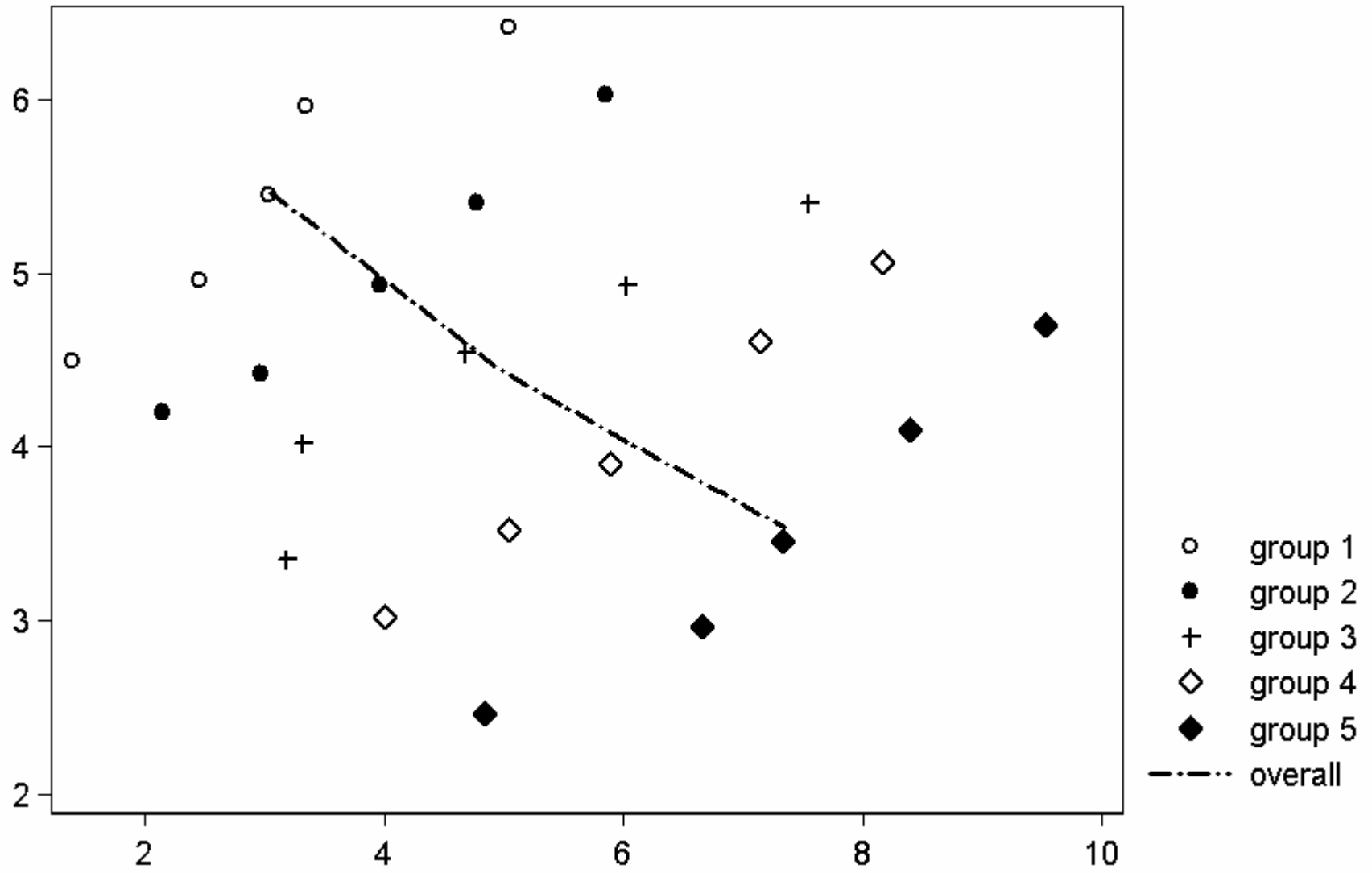
- Level-1 variables, such as student's gender and age
- Level-2 variables, such as school type and size

| schid | minority | female | ses | mathach | size | schtype | meanses |
|-------|----------|--------|--------|---------|------|---------|---------|
| 1224 | 0 | 0 | -.878 | 10.557 | 842 | 0 | -.428 |
| 1224 | 0 | 0 | -.938 | .868 | 842 | 0 | -.428 |
| 1224 | 0 | 0 | -.548 | 8.296 | 842 | 0 | -.428 |
| 1224 | 1 | 0 | .142 | -1.688 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | .972 | 2.059 | 842 | 0 | -.428 |
| 1224 | 0 | 0 | .372 | 6.714 | 842 | 0 | -.428 |
| 1224 | 1 | 0 | -1.658 | .122 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | -1.068 | 2.381 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | -.248 | 16.336 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | -1.398 | 6.134 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | .752 | 23.584 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | .012 | 14.053 | 842 | 0 | -.428 |
| 1224 | 0 | 1 | -.418 | 2.183 | 842 | 0 | -.428 |
| 1288 | 0 | 1 | -.788 | 7.857 | 1855 | 0 | .128 |
| 1288 | 1 | 0 | -.328 | 10.171 | 1855 | 0 | .128 |
| 1288 | 0 | 0 | .472 | 15.699 | 1855 | 0 | .128 |
| 1288 | 0 | 1 | .352 | 22.919 | 1855 | 0 | .128 |
| 1288 | 1 | 1 | -1.468 | 10.664 | 1855 | 0 | .128 |
| 1288 | 0 | 1 | .202 | 13.543 | 1855 | 0 | .128 |
| 1288 | 0 | 0 | -.518 | 18.207 | 1855 | 0 | .128 |
| 1288 | 1 | 0 | -.158 | 5.552 | 1855 | 0 | .128 |
| 1288 | 0 | 0 | .042 | 7.416 | 1855 | 0 | .128 |
| 1288 | 0 | 0 | .682 | 18.792 | 1855 | 0 | .128 |
| 1288 | 0 | 0 | 1.262 | 1.575 | 1855 | 0 | .128 |
| 1288 | 0 | 1 | .152 | 3.534 | 1855 | 0 | .128 |
| 1288 | 0 | 1 | -.678 | 20.173 | 1855 | 0 | .128 |
| 1288 | 0 | 1 | .332 | 10.772 | 1855 | 0 | .128 |

How would we analyze such multilevel data?

- OLS regression
- OLS regression with robust standard error
- Aggregation
- Disaggregation
- Ecological fallacy – interpreting analyses on aggregated data at the individual level

Ecological Fallacy



See figure 3.1, on page 14 from *Multilevel Analysis* by Snijders and Bosker

Hierarchical linear model

- Random Intercept model

$$Y_{ij} = \beta_{0j} + r_{ij}$$

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

- Written in mixed model format:

$$Y_{ij} = \gamma_{00} + u_{0j} + r_{ij}$$

- i is for individuals and j is for schools
- β_{0j} is the mean of Y_{ij} for school j
- γ_{00} is the average of all the β_{0j} 's, therefore the grand
- r_{ij} and u_{0j} are normally distributed
- r_{ij} and u_{0j} are independent of each other
- Parameters to be estimated include regression coefficients and variance components: γ_{00} , $\text{var}(r_{ij})$ and $\text{var}(u_{0j})$

Hierarchical linear model

- Random Intercept and random slope model

$$Y_{ij} = \beta_{0j} + \beta_{1j}X + r_{ij}$$

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + u_{1j}$$

- Written in mixed model format:

$$Y_{ij} = \gamma_{00} + \gamma_{10}X + u_{0j} + u_{1j}X + r_{ij}$$

- β_{0j} is the mean of Y_{ij} for school j when X is zero
- β_{1j} is the slope of X for school j (or the effect of X for school j)
- r_{ij} , u_{0j} and u_{1j} are normally distributed
- u_{0j} and u_{1j} are assumed to be correlated
- cross-level error terms are assumed to be independent
- parameters: γ_{00} , γ_{10} , $\text{var}(u_{0j})$, $\text{var}(u_{1j})$, $\text{cov}(u_{0j}, u_{1j})$ and $\text{var}(r_{ij})$

Hierarchical linear model

- Random Intercept and random slope model
- Level-2 variable(s) to predict intercept and/or slope

$$Y_{ij} = \beta_{0j} + \beta_{1j}X + r_{ij}$$

$$\beta_{0j} = \gamma_{00} + \gamma_{01}W + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}W + u_{1j}$$

- Written in mixed model format:

$$Y_{ij} = \gamma_{00} + \gamma_{01}W + \gamma_{10}X + \gamma_{11}W*X + u_{0j} + u_{1j}*X + r_{ij}$$

- β_{0j} is the mean of Y_{ij} for school j when X is zero
- β_{1j} is the slope of X for school j (or the effect of X for school j)
- γ_{00} is the average intercept
- γ_{11} is the coefficient for the cross-level interaction term
- r_{ij} , u_{0j} and u_{1j} are normally distributed
- u_{0j} and u_{1j} are assumed to be correlated
- Cross-level error terms are assumed to be independent
- parameters to be estimated: γ_{00} , γ_{01} , γ_{10} , γ_{11} , $\text{var}(u_{0j})$, $\text{var}(u_{1j})$, $\text{cov}(u_{0j}, u_{1j})$ and $\text{var}(r_{ij})$

Comparing the assumptions for hierarchical linear models with OLS models

OLS Assumptions

- Linearity: function form is linear
- Normality: residuals are normally distributed
- Homoscedasticity: residual variance is constant
- Independence: observations are independent of each other

HLM assumptions

- Linearity: function forms are linear at each level
- Normality: level-1 residuals are normally distributed and level-2 random effects u 's have a multivariate normal distribution
- Homoscedasticity: level-1 residual variance is constant
- Independence: level-1 residuals and level-2 residuals are uncorrelated
- Independence: observations at highest level are independent of each other

Estimation Methods: REML vs. ML

- Reading: Section 4.6 Parameter Estimation from Snijder and Bosker
- REML and ML produce similar regression coefficients
- REML and ML differ in terms of estimating the variance components
- If the number of level-2 units is small , then ML variance estimates will be smaller than REML, leading to artificially short confidence interval and biased significant tests.
- REML is the default estimation method for HLM
- Likelihood ratio test for nested models
 - When fixed effects are the same, model has fewer random effects , then both REML or ML may be used
 - When one model has fewer fixed effects and possibly fewer random effects, then ML may be used

Issues with Centering

- Reading: Section 5.2 *The effects of centering* from Kreft and De Leeuw
- In OLS centering is to change the interpretation of the intercept
- Centering in HLM is not a simple issue
- Grand-mean centering
“The raw score model and the grand mean centered model are *equivalent linear models*.”
- Group-mean centering
Most of the times, the group mean centered model and the raw score model are neither equivalent in the fixed part nor in the random part.
- Combining substantive and statistical reasons in choosing
 - raw score
 - group-centering with reintroducing the means
 - group-centering without reintroducing the means

An Example

- The dataset is a subsample from the 1982 High School and Beyond Survey and is used extensively in *Hierarchical Linear Models* by Raudenbush and Bryk.
- It consists of 7185 students nested in 160 schools.
- The outcome variable of interest is the student-level math achievement score, **mathach**.
- Predictor variables
 - Level-1 (student level) predictor variables:
 - **ses**: social-economic-status of a student
 - **female** 0 = male and 1 = female
 - Level-2 (school level) predictor variables:
 - **meanses**: mean ses at school level, aggregated from student level
 - **sctype**: type of school: 0 = public and 1 = private, there are 90 public and 70 private schools
 - **size**: size of a school

Model Building

- Reading: Section 6.4 Model specification from Snijder and Bosker

- Unconditional model:

$$\begin{aligned} \text{math}_{ij} &= \beta_{0j} + r_{ij} \\ \beta_{0j} &= \gamma_{00} + u_{0j} \end{aligned}$$

- Random intercept model with level-2 predictor(s):

$$\begin{aligned} \text{math}_{ij} &= \beta_{0j} + r_{ij} \\ \beta_{0j} &= \gamma_{00} + \gamma_{01}(\text{meanses}) + u_{0j} \end{aligned}$$

- Random intercept and random slope model:

$$\begin{aligned} \text{math}_{ij} &= \beta_{0j} + \beta_{1j}(\text{ses}) + r_{ij} \\ \beta_{0j} &= \gamma_{00} + u_{0j} \\ \beta_{1j} &= \gamma_{10} + u_{1j} \end{aligned}$$

- Full model:

$$\begin{aligned} \text{math}_{ij} &= \beta_{0j} + \beta_{1j}(\text{group_mean_centered_ses}) + r_{ij} \\ \beta_{0j} &= \gamma_{00} + \gamma_{01}(\text{sctype}) + \gamma_{02}(\text{meanses}) + u_{0j} \\ \beta_{1j} &= \gamma_{10} + \gamma_{11}(\text{sctype}) + \gamma_{12}(\text{meanses}) + u_{1j} \end{aligned}$$

Model 1: Unconditional Means Model

$$\text{mathachij} = \beta_{0j} + r_{ij} \quad \beta_{0j} = \gamma_{00} + u_{0j}$$

$$\gamma_{00} = 12.636972$$

$$\text{var}(r_{ij}) = 39.14831 \quad \text{var}(u_{0j}) = 8.61431$$

$$\begin{aligned} \text{Rho} &= \text{var}(u_{0j}) / (\text{var}(u_{0j}) + \text{var}(r_{ij})) \\ &= 8.61431 / (8.61431 + 39.14831) = .18035673 \end{aligned}$$

Final estimation of fixed effects
(with robust standard errors)

| Fixed Effect | Coefficient | Standard Error | T-ratio | Approx. d.f. | P-value |
|-----------------------------------|-------------|----------------|---------|--------------|---------|
| For INTRCPT1, B0 INTRCPT2, G00 | 12.636972 | 0.243628 | 51.870 | 159 | 0.000 |

Final estimation of variance components:

| Random Effect | Standard Deviation | Variance Component | df | Chi-square | P-value |
|----------------------------|--------------------|---------------------|-----|------------|---------|
| INTRCPT1, U0 level-1, R | 2.93501 6.25686 | 8.61431 39.14831 | 159 | 1660.23259 | 0.000 |

Final model

$$\text{mathachij} = \beta_{0j} + \beta_{1j}(\text{group_mean_centered_ses}) + r_{ij}$$

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{schtype}) + \gamma_{02}(\text{meanses}) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{schtype}) + \gamma_{12}(\text{meanses}) + u_{1j}$$

Tau

| | | |
|--------------|---------|---------|
| INTRCPT1, B0 | 2.37996 | 0.19058 |
| SES, B1 | 0.19058 | 0.14892 |

Final estimation of fixed effects
(with robust standard errors)

| Fixed Effect | Coefficient | Standard Error | T-ratio | Approx. d.f. | P-value |
|-------------------|-------------|----------------|---------|--------------|---------|
| ----- | | | | | |
| For INTRCPT1, B0 | | | | | |
| INTRCPT2, G00 | 12.096006 | 0.173699 | 69.638 | 157 | 0.000 |
| SCHTYPE, G01 | 1.226384 | 0.308484 | 3.976 | 157 | 0.000 |
| MEANSES, G02 | 5.333056 | 0.334600 | 15.939 | 157 | 0.000 |
| For SES slope, B1 | | | | | |
| INTRCPT2, G10 | 2.937981 | 0.147620 | 19.902 | 157 | 0.000 |
| SCHTYPE, G11 | -1.640954 | 0.237401 | -6.912 | 157 | 0.000 |
| MEANSES, G12 | 1.034427 | 0.332785 | 3.108 | 157 | 0.003 |

Final estimation of variance components:

| Random Effect | Standard Deviation | Variance Component | df | Chi-square | P-value |
|---------------|--------------------|--------------------|-----|------------|---------|
| ----- | | | | | |
| INTRCPT1, U0 | 1.54271 | 2.37996 | 157 | 605.29503 | 0.000 |
| SES slope, U1 | 0.38590 | 0.14892 | 157 | 162.30867 | 0.369 |
| level-1, R | 6.05831 | 36.70313 | | | |

Final Model (continued)

$$\begin{aligned} \text{mathach}_{ij} &= \beta_{0j} + \beta_{1j}(\text{group_mean_centered_ses}) + r_{ij} \\ \beta_{0j} &= \gamma_{00} + \gamma_{01}(\text{schtype}) + \gamma_{02}(\text{meanses}) + u_{0j} \\ \beta_{1j} &= \gamma_{10} + \gamma_{11}(\text{schtype}) + \gamma_{12}(\text{meanses}) + u_{1j} \end{aligned}$$

$\gamma_{00} = 12.096$: the intercept for public schools with meanses = 0 (average ses)

$\gamma_{01} = 1.226$: the change in intercept from a public school to a private school

– the intercept for private school with meanses = 0 is $12.096 + 1.226 = 13.322$

$\gamma_{02} = 5.333$: the change in intercept for a one-unit change in meanses

– the intercept for public school with meanses = 1 is $12.096 + 5.333 = 17.429$

$\gamma_{10} = 2.94$: the slope of gcses for public schools with meanses = 0.

– the effect of gcses for public schools with meanses = 0 is 2.94

$\gamma_{11} = -1.641$: the change in slope from a public school to a private school

– the effect of gcses for private schools with meanses = 0 is $2.94 - 1.641 = 1.299$

$\gamma_{12} = 1.034$: the change in slope for a one-unit change in meanses

– the effect of gcses for public schools with meanses = 0 is 2.94

– the effect of gcses for public schools with meanses = 1 is $2.94 + 1.034 = 3.974$

| | | | | | | |
|-----|---------------|-----------|----------|--------|-----|-------|
| For | INTRCPT1, B0 | | | | | |
| | INTRCPT2, G00 | 12.096006 | 0.173699 | 69.638 | 157 | 0.000 |
| | SCHTYPE, G01 | 1.226384 | 0.308484 | 3.976 | 157 | 0.000 |
| | MEANSES, G02 | 5.333056 | 0.334600 | 15.939 | 157 | 0.000 |
| For | SES slope, B1 | | | | | |
| | INTRCPT2, G10 | 2.937981 | 0.147620 | 19.902 | 157 | 0.000 |
| | SCHTYPE, G11 | -1.640954 | 0.237401 | -6.912 | 157 | 0.000 |
| | MEANSES, G12 | 1.034427 | 0.332785 | 3.108 | 157 | 0.003 |

What's new in HLM 6

The following paragraph is based on:

<http://www.ssicentral.com/hlm/new.html>

HLM 6 greatly broadens the range of hierarchical models that can be estimated. It also offers greater convenience of use than previous versions. Here is a quick overview of key new features and options:

- [All new graphical displays of data.](#)
- [Greater expanded graphics for fitted models.](#)
- Model equations displayed in hierarchical or mixed-model format with or without subscripts - easy to save for publication. Distribution assumptions and link functions are presented in detail.
- Slightly different and easier way for specifying random effects.
- [Cross-classified random effects models for linear models and non-linear link functions with convenient Windows interface.](#)
- High-order Laplace approximation with EM algorithm for stable convergence and accurate estimation in two-level hierarchical generalized linear models (HGLM).
- [Multinomial and ordinal models for three-level data. Also see the types of models.](#)
- [New flexible and accurate sample design weighting for two- and three-level HLMs and HGLMs.](#)
- Easier automated input from a wide variety of software packages, including the current versions of SAS, SPSS, and STATA.
- [Residual files can be saved directly as SPSS \(*.sav\) or STATA \(*.dta\) files.](#)
- [Analyses are based on MDM files, replacing the older less flexible SSM format.](#)

Getting ready for using HLM software for multilevel data analysis

- Creating MDM file
 - separate level-1 and level2 files for HLM2, or a single file
 - original file can be in different format, such as SPSS, Stata and SAS
 - linking variable can be either numeric or character
 - variables in the analyses have to be numeric
 - mdm file: binary file used for analyses and graphics
 - mdmt file: template file in text format for creating mdm file
 - hlm2mdm.sts: text file containing the summary statistics
- Data management
 - HLM does not have data management capability
 - One has to use other stat package(s) to clean the data and to create variables, such as dummy variables and within-level interaction terms
 - HLM handles cross-level interactions nicely

Choosing preferences and other settings

Preferences

type of non-ASCII data

SAS

SPSS

Stata

SYSTAT

other non-ASCII

Colors

Choose foreground color

Choose background color

OK

Show Mixed Model

Create Graph Files

Use level subscripts

Basic Model Specifications - HLM2

Distribution of Outcome Variable

Normal (Continuous)

Bernoulli (0 or 1)

Poisson (constant exposure)

Binomial (number of trials)

Poisson (variable exposure)

Multinomial

Ordinal

Number of categories

Over dispersion

Level-1 Residual File

Level-2 Residual File

Title: no title

Output file name: D:\work\test\hlm2.txt

Graph file name

Cancel OK

Estimation Settings - HLM2

Type of Likelihood

Restricted maximum likelihood

Full maximum likelihood

LaPlace Iteration Control

Do EM Laplace iterations

Maximum number of iterations

Constraint of fixed effects: Heterogeneous σ^2 Plausible values Multiple imputation

Level-1 Deletion Variables Weighting Latent Variable Regression

Fix σ^2 to specific value: computed

(Set to "computed" if you want σ^2 random or if over-dispersion is desired)

OK

Demo on using HLM

- Input Data and Creating the "MDM" file
 - from a single SPSS file
- data-based graphs
 - box-plot
 - scatter plot
- Model Building
 - unconditional means model
 - regression with means-as-outcomes
 - random-coefficient model
 - intercepts and slopes-as-outcomes model
- Hypothesis Testing, Model Fit
 - Multivariate hypothesis tests on fixed effects
 - Multivariate Tests of variance-covariance components specification
 - Model-based graphs
- Other Issues
 - Modeling Heterogeneity of Level-1 Variances
 - Models Without a Level-1 Intercept
 - Constraints on Fixed Effects

References

- [Multilevel Analysis: An Introduction to Basic and Advanced Multilevel Modeling](#) by Tom Snijders and Roel Bosker
- [Introduction to Multilevel Modeling](#) by Ita Kreft and Jan de Leeuw
- [Multilevel Analysis: Techniques and Applications](#) by Joop Hox
- Hierarchical Linear Models, Second Edition by Stephen Raudenbush and Anthony Bryk
- HLM 6 - Hierarchical Linear and Nonlinear Modeling by Raudenbush et al.