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Learning objectives and outline

GLM review

Interpretation of main effects and interactions in logistic regression

The Design Matrix

Session 3: Regression coefficients and model matrices

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CUNY SPH Biostatistics 2

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Learning objectives

- 1 Interpret main effect coefficients in logistic regression
- 2 Interpret interaction terms in logistic regression
- 3 Define and interpret model matrices for (generalized) linear models

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- 1 Review of GLM
- 2 Interpretation of logistic regression coefficients
- 3 Introduction to model matrices

Outline

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Components of GLM

- **Random component** specifies the conditional distribution for the response variable
 - doesn't have to be normal
 - can be any distribution in the "exponential" family of distributions
- Systematic component specifies linear function of predictors (linear predictor)
- Link [denoted by g(.)] specifies the relationship between the expected value of the random component and the systematic component
 - can be linear or nonlinear

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Logistic Regression as GLM

• The model:

$$Logit(P(x)) = log\left(\frac{P(x)}{1 - P(x)}\right)$$
$$= \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \beta_p x_{pi}$$

- **Random component**: *y_i* follows a Binomial distribution (outcome is a binary variable)
 - Systematic component: linear predictor

 $\beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \ldots + \beta_p x_{pi}$

• Link function: *logit* (log of the odds that the event occurs)

$$g(P(x)) = logit(P(x)) = log\left(rac{P(x)}{1-P(x)}
ight)$$

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Additive vs. multiplicative models

- 1 Linear regression is an *additive* model
 - e.g. for two binary variables $\beta_1 = 1.5$, $\beta_2 = 1.5$.
 - If $x_1 = 1$ and $x_2 = 1$, this adds 3.0 to E(y|x)
- 2 Logistic regression is a *multiplicative* model
 - It is additive on *log*-odds scale
 - If $x_1 = 1$ and $x_2 = 1$, this adds 3.0 to $log(\frac{P}{1-P})$
 - Odds-ratio $\frac{P}{1-P}$ increases 20-fold: exp(1.5 + 1.5) or exp(1.5) * exp(1.5)

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Motivating example: contraceptive use data

From http://data.princeton.edu/wws509/datasets/#cuse

cuse <- read.table("cuse.dat", header=TRUE)
summary(cuse)</pre>

##	age	education	wantsMore	notUsing
##	Length:16	Length:16	Length:16	Min. : 8.00
##	Class :character	Class :character	Class :character	1st Qu.: 31.00
##	Mode :character	Mode :character	Mode :character	Median : 56.50
##				Mean : 68.75
##				3rd Qu.: 85.75
##				Max. :212.00
##	using			
##	Min. : 4.00			
##	1st Qu.: 9.50			
##	Median :29.00			
##	Mean :31.69			
##	3rd Qu.:49.00			
##	Max. :80.00			

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Univariate regression on "wants more children"

```
fit <- glm(cbind(using, notUsing) ~ wantsMore,</pre>
          data=cuse. family=binomial("logit"))
summarv(fit)
##
## Call:
## glm(formula = cbind(using, notUsing) ~ wantsMore, family = binomial("logit"),
##
      data = cuse)
##
## Coefficients:
##
               Estimate Std. Error z value Pr(>|z|)
## (Intercept) -0.18636 0.07971 -2.338 0.0194 *
## wantsMoreyes -1.04863 0.11067 -9.475 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##
       Null deviance: 165.772 on 15 degrees of freedom
## Residual deviance: 74.098 on 14 degrees of freedom
## ATC 149.61
##
## Number of Fisher Scoring iterations: 4
```

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Interpretation of "wants more children" table

- Coefficients for (Intercept) and dummy variables
- Coefficients are normally distributed when assumptions are correct

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Interpretation of "wants more children" coefficients

Additive coefficient interpretation on log-odds scale



Wants more children

Figure 1: Diagram of the estimated coefficients in the GLM. The yellow arrow indicates the Intercept term, which goes from zero to the mean of the reference group (here the 'wantsMore = no' samples). The blue arrow indicates the difference in log-odds of the yes group minus the no group, which is negative in this example. The circles show the individual samples, jittered horizontally to avoid

Regression on age

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- Four age groups
 - three dummy variables age25-29, age30-39, age40-49
 - how to interpret them?

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Regression on age

```
data=cuse. family=binomial("logit"))
summary(fit)
##
## Call:
## glm(formula = cbind(using, notUsing) ~ age, family = binomial("logit").
      data = cuse)
##
##
## Coefficients:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) -1.5072 0.1303 -11.571 < 2e-16 ***
## age25-29 0.4607 0.1727 2.667 0.00765 **
## age30-39 1.0483 0.1544 6.788 1.14e-11 ***
## age40-49 1.4246 0.1940 7.345 2.06e-13 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 165.772 on 15 degrees of freedom
## Residual deviance: 86.581 on 12 degrees of freedom
## AIC: 166.09
##
## Number of Fisher Scoring iterations: 4
```

fit <- glm(cbind(using, notUsing) ~ age,

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Recall model formulae

and	symbol	example	meaning
ew ntion	+	+ x	include this variable
	-	- X	delete this variable
d ns in	:	x : z	include the interaction
	*	× * 7	include these variables and their
gn		~ <u>-</u>	interactions
	^	$(u + v + w)^{3}$	include these variables and all
			interactions up to three way
	1	-1	intercept: delete the intercept

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Regression on age and wantsMore

	Estimate	Std. Error	z value	$\Pr(> z)$
(Intercept)	-0.87	0.16	-5.54	0.00
age25-29	0.37	0.18	2.10	0.04
age30-39	0.81	0.16	5.06	0.00
age40-49	1.02	0.20	5.01	0.00
wantsMoreyes	-0.82	0.12	-7.04	0.00

Interaction / Effect Modification

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The Design Matrix What if we want to know whether the effect of age is modified by whether the woman wants more children or not?

Interaction is modeled as the product of two covariates:

 $E[y|x] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_{12} x_1 * x_2$

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Interaction / Effect Modification (fit)

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.46	0.30	-4.90	0.00
age25-29	0.64	0.36	1.78	0.07
age30-39	1.54	0.32	4.84	0.00
age40-49	1.76	0.34	5.14	0.00
wantsMoreyes	-0.06	0.33	-0.19	0.85
age25-29:wantsMoreves	-0.27	0.41	-0.65	0.51
age30-39:wantsMoreyes	-1.09	0.37	-2.92	0.00
age40-49:wantsMoreyes	-1.37	0.48	-2.83	0.00

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What is the design matrix, and why?

- **1** What? The design matrix is the most generic, flexible way to specify them
- 2 Why? There are multiple possible and reasonable regression models for a given study design.

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The Design Matrix

or simply:

Matrix notation for the multiple linear regression model

$$\begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_N \end{pmatrix} = \begin{pmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots \\ 1 & x_N \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_N \end{pmatrix}$$

$$\mathbf{Y} = \mathbf{X}oldsymbol{eta} + oldsymbol{arepsilon}$$

- The design matrix is X
- the computer will take X as a given when solving for β by minimizing the sum of squares of residuals ε, or maximizing likelihood.

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Choice of design matrix

• The model formula encodes a default model matrix, e.g.:

```
group <- factor( c(1, 1, 2, 2) )
model.matrix(~ group)</pre>
```

##	(Intercept)	group2
##	1 1	0
##	2 1	0
##	3 1	1
##	4 1	1
##	attr(,"assign	")
##	[1] 0 1	
##	attr(,"contra	sts")
##	attr(,"contra	sts")\$group
##	[1] "contr.tr	eatment"

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The Design Matrix

Choice of design matrix (cont'd)

What if we forgot to code group as a factor?

```
group <- c(1, 1, 2, 2)
model.matrix(~ group)</pre>
```

##	(I	ntercept)	group
##	1	1	1
##	2	1	1
##	3	1	2
##	4	1	2
##	attr	(,"assign	")
##	[1]	0 1	

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More groups, still one variable

group <- factor(c(1,1,2,2,3,3))
model.matrix(~ group)</pre>

##	(Intercept	t)	group2	group3
##	1	1	0	0
##	2	1	0	0
##	3	1	1	0
##	4	1	1	0
##	5	1	0	1
##	6	1	0	1
##	attr(,"assig	gn'	')	
##	[1] 0 1 1			
##	attr(,"cont	ras	sts")	
##	attr(,"cont	ras	sts")\$gi	coup
##	[1] "contr.	tre	eatment'	1

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Changing the baseline group

```
group <- factor(c(1,1,2,2,3,3))
group <- relevel(x=group, ref=3)
model.matrix(~ group)</pre>
```

##	(Intercept)	group1	group2			
##	1 1	1	0			
##	2 1	1	0			
##	3 1	0	1			
##	4 1	0	1			
##	5 1	0	0			
##	6 1	0	0			
##	attr(,"assign")					
##	[1] 0 1 1					
##	attr(,"contra	sts")				
##	attr(,"contra	sts")\$gi	coup			
##	[1] "contr.tr	eatment'	1			

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More than one variable

```
agegroup <- factor(c(1,1,1,1,2,2,2,2))
wantsMore <- factor(c("y","y","n","n","y","y","n","n"))</pre>
model.matrix(~ agegroup + wantsMore)
```

##	(Intercept)	agegroup2	wantsMorey
##	1 1	0	1
##	2 1	0	1
##	3 1	0	0
##	4 1	0	0
##	5 1	1	1
##	6 1	1	1
##	7 1	1	0
##	8 1	1	0
##	attr(,"assign	")	
##	[1] 0 1 2		
##	attr(,"contra	sts")	
##	attr(,"contra	sts")\$agegi	roup
##	[1] "contr.tr	eatment"	-
##			
##	attr(,"contra	sts")\$want:	sMore
##	[1] "contr.tr	eatment"	

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With an interaction term

	<i>′</i> -			
##	(Intercept)	agegroup2	wantsMorey	agegroup2:wantsMor
## 1	1	0	1	
## 2	1	0	1	
## 3	1	0	0	
## 4	1	0	0	
## 5	1	1	1	
## 6	1	1	1	
## 7	1	1	0	
## 8	1	1	0	
## a	ttr(,"assign	")		
## [1] 0 1 2 3			
## a	ttr(,"contra	sts")		
## a	ttr(,"contra	sts")\$agegi	roup	
## [1] "contr.tr	eatment"	-	
##				
## a	ttr(."contra	sts")\$wants	sMore	
## Г	1] "contr tr	ootmont"		

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The Design Matrix

Design matrix to contrast what we want

- Contraceptive use example
 - The effect of wanting more children different for 40-49 year-olds than for <25 year-olds is answered by the term age40-49:wantsMoreyes in this default model with interaction terms

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.46	0.30	-4.90	0.00
age25-29	0.64	0.36	1.78	0.07
age30-39	1.54	0.32	4.84	0.00
age40-49	1.76	0.34	5.14	0.00
wantsMoreyes	-0.06	0.33	-0.19	0.85
age25-29:wantsMoreyes	-0.27	0.41	-0.65	0.51
age30-39:wantsMoreyes	-1.09	0.37	-2.92	0.00
age40-49:wantsMoreyes	-1.37	0.48	-2.83	0.00

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Design matrix to contrast what we want (cont'd)

• What if we want to ask this question for 40-49 year-olds vs. 30-39 year-olds?

The desired contrast is:

age40-49:wantsMoreyes - age30-39:wantsMoreyes

There are many ways to construct this design, one is with library(multcomp)

Design matrix constructed with library(multcomp)

coef(fit) (Intercept) age25-29 age30-39 ## -1.455287230.63538835 1.54114852 ## age40-49 wantsMoreyes age25-29:wantsMoreyes ## ## 1 76429207 -0.06399958 -0.26723185 ## age30-39:wantsMoreyes age40-49:wantsMoreyes -1.09049316-1.36714805## contmat <- matrix(c(0,0,0,0,0,0,-1,1), 1)contmat ## [.1] [.2] [.3] [.4] [.5] [.6] [.7] [.8] ## [1.] 0 0 0 0 0 0 -1 1 new.interaction <- multcomp::glht(fit, linfct=contmat) summary(new.interaction) ## Simultaneous Tests for General Linear Hypotheses ## ## ## Fit: glm(formula = cbind(using, notUsing) ~ age * wantsMore, family = binomial("logit"), data = cuse) ##

```
## Linear Hypotheses:
## Estimate Std. Error z value Pr(>|z|)
## 1 == 0 -0.2767 0.3935 -0.703 0.482
## (Adjusted p values reported -- single-step method)
```

Session 3: Regression coefficients and model matrices

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Summary

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Session 3: Regression

coefficients and model

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The Design Matrix

- Logistic regression coefficients are *linear* in log-odds, *multiplicative* in probability
- 2 model formulae for easy setup of multiple regression
- 3 design matrix for completely flexible setup of multiple regression