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Session 10: Repeated Measures and Longitudinal Analysis II

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CUNY SPH Biostatistics 2

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Learning objectives

- **1** Define mixed effects models and population average models
- **2** Perform model diagnostics for random effects models
- **3** Interpret random intercepts and random slopes
- **4** Define and perform population average models
- **5** Define assumptions on correlation structure in hierarchical models
- **6** Choose between hierarchical modeling strategies

Outline

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1 Review of fecal fat dataset

- **2** Summary of non-hierarchical approaches
- **3** Mixed effects models
- **4** Longitudinal data and the Georgia Birthweights dataset
- **5** Population average models and Generalized Estimating Equations (GEE)
- Vittinghoff sections 7.2, 7.3, 7.5

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Fecal fat dataset

- Lack of digestive enzymes in the intestine can cause bowel absorption problems.
	- This will be indicated by excess fat in the feces.
	- Pancreatic enzyme supplements can alleviate the problem.
	- fecfat.csv: a study of fecal fat quantity (g/day) for individuals given each of a placebo and 3 types of pills

Table 7.1 Fecal fat (g/day) for six subjects

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Fecal fat dataset

Warning: Using 'size' aesthetic for lines was depr ## i Please use 'linewidth' instead.

This warning is displayed once every 8 hours.

Call 'lifecycle::last_lifecycle_warnings()' to see ## generated.

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Analysis strategies for hierarchical data

- Fixed effects and other non-hierarchical strategies
- Random / mixed effects models
	- model certain regression coefficients (intercept, slopes) as random variables
- Population average models
	- using Generalized Estimating Equations (GEE)

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Non-hierarchical analysis strategies for hierarchical data

- Analyses for each subgroup
	- e.g., look at each patient independently
	- doesn't work at all in this example, and in general is not an integrated analysis of the whole data
	- could sort of work for an example with many patients per doctor, a few doctors
- Analysis at the highest level in the hierarchy
	- first summarize data to highest level
	- doesn't work at all in this example
	- could sort of work for an example with few patients per doctor, many doctors
- Analysis on "Derived Variables"
	- consider each treatment type separately, take differences in fat levels between treatment/control for each patient and use paired t-tests
	- can work, but not for unbalanced groups
- Fixed-effects models

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When is hierarchical analysis definitely needed?

- **1** the correlation structure is of interest, e.g. familial aggregation of disease, or consistency of treatment within centers
- **2** we wish to "borrow strength" across the levels of a hierarchy in order to improve estimates
- **3** dealing with unbalanced data
- **4** we want to benefit from software designed for hierarchical data

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Mixed effects models

• Model looks like two-way ANOVA:

 $FECFAT_{ii} = \beta_0 + \beta_{subject} SUBJECT_i + \beta_{billtvoei} PILLT YPE_i + \epsilon_{ii}$

● Assumption: $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma_{\epsilon}^2)$

• But instead of fitting a *β* to each individual, we assume that the subject effects are selected from a distribution of possible subject effects:

 $FECFAT_{ii} = \beta_0 + SUBJECT_i + \beta_{pilltvpei} PILLTYPE_i + \epsilon_{ii}$

Where we assume: $\mathit{SUBJECT}_i \stackrel{iid}{\sim} N(0, \tau_{00}^2)$

- This is a *mixed effects* model because:
	- the "true" intercept varies randomly from patient to patient
	- the "true" (population) coefficient of treatment is fixed (the same for everyone)

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Fit this mixed-effects model

```
library(nlme)
fitmix <- nlme::lme(fecfat ~ pilltype,
                    data = dat,random = \sim 1 | subject)
```
Note: the lme4 package is another popular alternative

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```
Mixed effects
models
```
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Mixed effects model coeffients, variances, ICC

```
## Linear mixed-effects model fit by REML
    Data: dat
## Log-restricted-likelihood: -84.55594
## Fixed: fecfat ~ pilltype
## (Intercept) pilltypetablet pilltypecapsule pilltypecoated
## 38.083334 -21.550001 -20.666667 -7.016668
##
## Random effects:
## Formula: ~1 | subject
## (Intercept) Residual
            15.89557 10.34403
##
## Number of Observations: 24
## Number of Groups: 6
```
 $\text{ICC} = 15.9^2 / (15.9^2 + 10.34^2) = 0.7 = 0.7.$

- Recall ICC is a measure of how large the subject effect is, in relation to the error term
- Variances were estimated directly by the model!

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Assumptions of the mixed model

 $FECFAT_{ii} = \beta_0 + SUBJECT_i + \beta_{pilltvoei} PILLTYPE_i + \epsilon_{ii}$

- Normally distributed residuals as in fixed effects model:
	- \bullet *∈_i* $\stackrel{iid}{\sim} N(0, \sigma_{\epsilon}^2)$
- Normally distributed **latent variable**:
	- *SUBJECT_i* ^{iid} N(0, τ₀₀)

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Mixed effects model results

A plot of the random intercept:

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 $^{\prime}$

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Mixed effects model diagnostics

QQ plot residuals

QQ plot random intercepts

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Mixed effects model results

```
## Linear mixed-effects model fit by REML
## Data: dat
## AIC BIC logLik<br>## 181 1119 187 0863 -84 55594
    181.1119 187.0863 -84.55594
##
## Random effects:
## Formula: ~1 | subject
## (Intercept) Residual<br>## StdDev: 15,89557,10,34403
            15.89557 10.34403
##
## Fixed effects: fecfat ~ pilltype
## Value Std.Error DF t-value p-value
                  38.08333 7.742396 15 4.918805 0.0002
## pilltypetablet -21.55000 5.972127 15 -3.608430 0.0026
## pilltypecapsule -20.66667 5.972127 15 -3.460521 0.0035
## pilltypecoated -7.01667 5.972127 15 -1.174903 0.2583
## Correlation:
## (Intr) plltypt plltypcp
## pilltypetablet -0.386
## pilltypecapsule -0.386 0.500
## pilltypecoated -0.386 0.500 0.500
##
## Standardized Within-Group Residuals:
## Min Q1 Med Q3 Max
## -1.210052934 -0.615068039 -0.002727166 0.457105344 1.725618643
##
## Number of Observations: 24
## Number of Groups: 6
```
- Note: correlation of the estimator of the fixed effects
	- high correlations may (but not necessarily) be due to collinearity

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Mixed effects model results

Inference for variance terms (and fixed effects):

```
## Approximate 95% confidence intervals
##
## Fixed effects:
## lower est. upper
## (Intercept) 21.58081 38.083334 54.585860
## pilltypetablet -34.27929 -21.550001 -8.820714
## pilltypecapsule -33.39595 -20.666667 -7.937381
## pilltypecoated -19.74595 -7.016668 5.712618
##
## Random Effects:<br>## Level: subject
    Level: subject
## lower est. upper
## sd((Intercept)) 8.00117 15.89557 31.57904
##
## Within-group standard error:
## lower est. upper
  7.23240 10.34403 14.79438
```
- Would conclude that variation of the intercept between subjects is non-zero
	- not attributable to within-subject variation

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- Interested in the change in the value of a variable within a "subject"
- Collect data repeatedly through time.
- For hierarchical longitudinal analysis to be effective, before/after measurements need to be positively correlated

Longitudinal data

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Longitudinal data examples

- Example 1: a measure of sleepiness before and after administration of treatment or placebo
- Example 2: Study of Osteoporotic Fractores (SOF dataset)
	- 9,704 women tracked with clinical visits every two years
	- Bone Mineral Density (BMD), Body Mass Index (BMI), many other variables
- Questions for Example 2:
	- **1** Is change in BMD related to age at menopause? This is a time-invariant predictor, age at menopause, with time-dependent changes in the outcome, BMD.
	- **2** Is change in BMD related to change in BMI? This is an analysis relating a time-varying predictor, BMI, with changes in the outcome, BMD. BMI varies quite a lot between women, but also varies within a woman over time.

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Longitudinal data examples

- birthweight and birth order
- provides birthweights and order of infants from mothers who had 5 children in Georgia
	- interested in whether birthweight of babies changes with order
	- whether this difference depends on the *mother's age at* first childbirth or on the weight of initial baby.

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Birth Weight (g)

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Georgia Birthweights dataset

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Georgia Birthweights dataset

• Does baseline birth weight vary by mother? • random intercept

```
library(nlme)
```

```
gafit1 <- lme(bweight ~ birthord, data=ga,
              random=~1|momid)
```
Note: there are not enough degrees of freedom to also fit a random coefficient for birth order

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momid

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Georgia Birthweights dataset

Random effects

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Georgia Birthweights dataset

summary(gafit1)

```
## Linear mixed-effects model fit by REML
## Data: ga
## AIC BIC logLik
    15321.65 15341.28 -7656.826
##
## Random effects:
## Formula: ~1 | momid
## (Intercept) Residual<br>## StdDev: 367,2676,445,0228
            367.2676 445.0228
##
## Fixed effects: bweight ~ birthord
                Value Std.Error DF t-value p-value
## (Intercept) 2995.640 41.99615 799 71.33130 0
               46.608 9.95101 799 4.68374
## Correlation:
## (Intr)
## birthord -0.711
##
## Standardized Within-Group Residuals:
## Min Q1 Med Q3 Max
## -5.26801358 -0.43683345 0.05028638 0.52703429 3.30770805
##
## Number of Observations: 1000
## Number of Groups: 200
```
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Georgia Birthweights dataset

intervals(gafit1, which = "all")

```
## Approximate 95% confidence intervals
##
## Fixed effects:
## lower est. upper
  (Intercept) 2913.20418 2995.640 3078.07582<br>hirthord 27.07478 46.608 66.14122
## birthord 27.07478 46.608 66.14122
##
## Random Effects:
## Level: momid
## lower est. upper
## sd((Intercept)) 323.1724 367.2676 417.3794
##
## Within-group standard error:
## lower est. upper
## 423.7298 445.0228 467.3859
```
- Does baseline birth weight vary by mother?
	- yes: the subject variance is significantly greater than zero
	- The variance between mothers is too much to be explained by within-mother variation in birth weights

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Population Average Models

- An alternative to random / mixed-effects models that is more robust to assumptions of:
	- distribution of random effects
	- correlation structure
- Estimates correlation structure from the data rather than assuming normality
	- Requires more clusters than observations per cluster
- Estimates regression coefficients and robust standard errors
	- commonly by Generalized Estimating Equations (GEE)

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Population Average Models

• Compare mixed model multiple linear regression:

 $E[Y_{ii} | X_{ii}] = \beta_0 + \alpha_{0i} + \beta_1 X_{ii}, \alpha_{0i} \sim N(0, \sigma)$

for subject i in group j .

• to a population average model:

$$
E[Y_{ij}|X_{ij}] = \beta_0^* + \beta_1^* X_{ij}
$$

- Interpretations of *β* ∗ and *β* are equivalent
- Numerically equivalent for linear and log-linear models (if specification of mixed model is correct), but not for logistic link.

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Fit a population average model

```
gafit.gee <- gee::gee(bweight ~ birthord,
                        corstr = "exchangeable",
                        id = \text{monid},
                        data = ga
```


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Correlation assumptions for GEE

Must make some assumption about the form of correlation among grouped observations. Some options are:

- Independence:
	- no correlation between measurements within group
- Exchangeable:
	- all pairwise correlations are the same (in large-N limit)
	- nothing distinguishes one member of a cluster from another
	- appropriate in the absence of other data structures such as measurements taken through time or space
- Auto-regressive $(AR-M)$:
	- observations taken more closely in time are more highly correlated

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Correlation assumptions for GEE (cont'd)

- Unstructured:
	- estimates a separate correlation between observations taken on each pair of "times"
- Non-stationary ("non_stat_M_dep"):
	- similar to unstructured, but assumes all correlations for pairs separated far enough in time are zero
- Stationary ("stat_M_dep"):
	- e.g. stationary of order 2: observations taken at time points 1 and 3 have the same correlation as time points 2 and 4
	- but this might be different from the correlation between observations taken at times 2 and 3
	- correlations for observations 3 or more time periods apart assumed to be zero

Fewer assumptions requires more data, and good assumptions improve results

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Help in choosing a method

Note: ^aOnly for calculation of standard errors. ^bProblems can arise under some specifications of the working covariance structure and depending on the estimation method used.

doi:10.1371/journal.pone.0146721.t002

Figure 2: Hierarchical modeling decision table from Moen et al.

Conclusions

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- Ignoring within-subject correlations can produce very wrong results, and is not always "conservative"
- Hierarchical analysis strategies are needed for any of:
	- **1** When the correlation structure is of primary interest, e.g. familial aggregation of disease, or consistency of treatment within centers,
	- **2** When we wish to "borrow strength" across the levels of a hierarchy in order to improve estimates, and
	- **3** When dealing with unbalanced correlated data. E.g., no requirement that each Georgia mother have exactly 5 children.
- Population average models provide a robust alternative to mixed models
	- for one level of hierarchy

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A final note on reporting results of hypothesis tests

- Include test statistic, a measure of "effect size", and test name if unclear from test statistic
- Write in plain language and let the statistics support, not lead. E.g.:
	- do: The 36 study participants had a mean age of 27.4 (SD $= 12.6$), significantly older than the university mean of 21.2 years $(t(35) = 2.95, p = 0.01)$.
	- don't: A p-value of 0.01 indicated significant difference in age of study participants compared to all university students.
	- do: report confidence intervals where possible
- [UW "Reporting Results of Common Statistical Tests in APA Format":](https://psych.uw.edu/storage/writing_center/stats.pdf) specific examples of reporting a hypothesis test result
- STROBE guidelines for reporting observational studies:<https://www.strobe-statement.org/>
- [A Guideline for Reporting Results of Statistical Analysis in Japanese Journal of Clinical Oncology:](https://academic.oup.com/DocumentLibrary/JJCO/eng.guideline.pdf) helpful guidelines for all parts of a manuscript

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CONGRATULATIONS!!!