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Learning objectives and outline

Review

Nonhierarchical analysis strategies

Mixed effects models

Longitudinal data

Population Average Models

Session 10: Repeated Measures and Longitudinal Analysis II

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CUNY SPH Biostatistics 2

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Learning objectives

- 1 Define mixed effects models and population average models
- 2 Perform model diagnostics for random effects models
- 3 Interpret random intercepts and random slopes
- 4 Define and perform population average models
- **5** Define assumptions on correlation structure in hierarchical models
- 6 Choose between hierarchical modeling strategies

Outline

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- 1 Review of fecal fat dataset
- 2 Summary of non-hierarchical approaches
- 3 Mixed effects models
- 4 Longitudinal data and the Georgia Birthweights dataset
- 5 Population average models and Generalized Estimating Equations (GEE)
- Vittinghoff sections 7.2, 7.3, 7.5

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Fecal fat dataset

- Lack of digestive enzymes in the intestine can cause bowel absorption problems.
 - This will be indicated by excess fat in the feces.
 - Pancreatic enzyme supplements can alleviate the problem.
 - fecfat.csv: a study of fecal fat quantity (g/day) for individuals given each of a placebo and 3 types of pills

Subject number	Pill type				Subject
	None	Tablet	Capsule	Coated	Average
1	44.5	7.3	3.4	12.4	16.9
2	33.0	21.0	23.1	25.4	25.6
3	19.1	5.0	11.8	22.0	14.5
4	9.4	4.6	4.6	5.8	6.1
5	71.3	23.3	25.6	68.2	47.1
6	51.2	38.0	36.0	52.6	44.5
Pill type					
average	38.1	16.5	17.4	31.1	25.8

Table 7.1 Fecal fat (g/day) for six subjects

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Fecal fat dataset

Warning: Using 'size' aesthetic for lines was depr ## i Please use 'linewidth' instead.

This warning is displayed once every 8 hours.

Call 'lifecycle::last_lifecycle_warnings()' to see
generated.



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Analysis strategies for hierarchical data

- Fixed effects and other non-hierarchical strategies
- Random / mixed effects models
 - model certain regression coefficients (intercept, slopes) as random variables
- Population average models
 - using Generalized Estimating Equations (GEE)

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Non-hierarchical analysis strategies

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Non-hierarchical analysis strategies for hierarchical data

- Analyses for each subgroup
 - e.g., look at each patient independently
 - doesn't work at all in this example, and in general is not an integrated analysis of the whole data
 - could sort of work for an example with many patients per doctor, a few doctors
- Analysis at the highest level in the hierarchy
 - first summarize data to highest level
 - doesn't work at all in this example
 - could sort of work for an example with few patients per doctor, many doctors
- Analysis on "Derived Variables"
 - consider each treatment type separately, take differences in fat levels between treatment/control for each patient and use paired t-tests
 - can work, but not for unbalanced groups
- Fixed-effects models

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When is hierarchical analysis definitely needed?

- 1 the correlation structure is of interest, *e.g.* familial aggregation of disease, or consistency of treatment within centers
- 2 we wish to "borrow strength" across the levels of a hierarchy in order to improve estimates
- 3 dealing with unbalanced data
- 4 we want to benefit from software designed for hierarchical data

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Mixed effects models

• Model looks like two-way ANOVA:

 $\textit{FECFAT}_{ij} = \beta_0 + \beta_{\textit{subjecti}} \textit{SUBJECT}_i + \beta_{\textit{pilltypej}} \textit{PILLTYPE}_j + \epsilon_{ij}$

• Assumption: $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma_{\epsilon}^2)$

 But instead of fitting a β to each individual, we assume that the subject effects are selected from a distribution of possible subject effects:

 $\textit{FECFAT}_{ij} = \beta_0 + \textit{SUBJECT}_i + \beta_{\textit{pilltypej}}\textit{PILLTYPE}_j + \epsilon_{ij}$

Where we assume: $SUBJECT_i \stackrel{iid}{\sim} N(0, \tau_{00}^2)$

- This is a *mixed effects* model because:
 - the "true" intercept varies randomly from patient to patient
 - the "true" (population) coefficient of treatment is fixed (the same for everyone)

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Fit this mixed-effects model

Note: the 1me4 package is another popular alternative

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```
Mixed effects 
models
```

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Mixed effects model coefficients, variances, ICC

```
## Linear mixed-effects model fit by REML
##
     Data: dat
##
     Log-restricted-likelihood: -84.55594
##
     Fixed: fecfat ~ pilltype
       (Intercept) pilltypetablet pilltypecapsule pilltypecoated
##
##
         38.083334
                        -21.550001
                                         -20.666667
                                                           -7.016668
##
## Bandom effects:
   Formula: ~1 | subject
##
##
           (Intercept) Residual
## StdDev:
              15 89557 10 34403
##
## Number of Observations: 24
## Number of Groups: 6
```

 $ICC = 15.9^2/(15.9^2 + 10.34^2) = 0.7 = 0.7.$

- Recall ICC is a measure of how large the subject effect is, in relation to the error term
- Variances were estimated directly by the model!

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Assumptions of the mixed model

$$FECFAT_{ij} = \beta_0 + SUBJECT_i + \beta_{pilltypej}PILLTYPE_j + \epsilon_{ij}$$

- Normally distributed residuals as in fixed effects model:
 - $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma_{\epsilon}^2)$

ŀ

- Normally distributed latent variable:
 - SUBJECT_i $\stackrel{iid}{\sim} N(0, \tau_{00}^2)$

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Mixed effects model results

A plot of the random intercept:



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Mixed effects model diagnostics

QQ plot residuals

QQ plot random intercepts



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Mixed effects model results

```
## Linear mixed-effects model fit by REML
##
     Data: dat
##
          ATC
                  BIC
                          logLik
    181,1119 187,0863 -84,55594
##
##
## Bandom effects:
  Formula: ~1 | subject
##
          (Intercept) Residual
##
## StdDev:
              15 89557 10 34403
##
## Fixed effects: fecfat ~ pilltype
##
                      Value Std.Error DF
                                           t-value p-value
## (Intercept)
                    38 08333 7 742396 15 4 918805
                                                    0.0002
## pilltypetablet -21.55000
                             5.972127 15 -3.608430
                                                    0.0026
## pilltypecapsule -20.66667
                             5.972127 15 -3.460521
                                                    0.0035
## pilltypecoated -7.01667
                              5.972127 15 -1.174903 0.2583
## Correlation:
##
                   (Intr) plltypt plltypcp
## pilltypetablet -0.386
## pilltypecapsule -0.386
                          0.500
## pilltypecoated -0.386 0.500
                                   0.500
##
## Standardized Within-Group Residuals:
##
           Min
                                      Med
                                                    03
                                                                Max
## -1.210052934 -0.615068039 -0.002727166
                                         0.457105344 1.725618643
##
## Number of Observations: 24
## Number of Groups: 6
```

- Note: correlation of the estimator of the fixed effects
 - high correlations may (but not necessarily) be due to collinearity

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Inference for variance terms (and fixed effects):

```
## Approximate 95% confidence intervals
##
   Fixed effects:
##
##
                       lower
                                    est.
                                             upper
## (Intercept)
                    21.58081
                               38.083334 54.585860
## pilltypetablet -34.27929 -21.550001 -8.820714
## pilltypecapsule -33.39595 -20.6666667 -7.937381
## pilltypecoated -19.74595 -7.016668 5.712618
##
   Random Effects:
##
##
    Level: subject
##
                     lower
                                est.
                                        upper
## sd((Intercept)) 8.00117 15.89557 31.57904
##
   Within-group standard error:
##
##
      lower
                est.
                        upper
   7 23240 10 34403 14 79438
##
```

- Would conclude that variation of the intercept between subjects is non-zero
 - not attributable to within-subject variation

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Population Average Models

- Interested in the change in the value of a variable within a "subject"
- Collect data repeatedly through time.
- For hierarchical longitudinal analysis to be effective, before/after measurements need to be positively correlated

Longitudinal data

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- Interested in the change in the value of a variable within a "subject"
- Collect data repeatedly through time.
- For hierarchical longitudinal analysis to be effective, before/after measurements need to be positively correlated

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Longitudinal data examples

- Example 1: a measure of sleepiness before and after administration of treatment or placebo
- Example 2: Study of Osteoporotic Fractores (SOF dataset)
 - 9,704 women tracked with clinical visits every two years
 - Bone Mineral Density (BMD), Body Mass Index (BMI), many other variables
- Questions for Example 2:
 - 1 Is change in BMD related to age at menopause? This is a time-invariant predictor, age at menopause, with time-dependent changes in the outcome, BMD.
 - 2 Is change in BMD related to change in BMI? This is an analysis relating a time-varying predictor, BMI, with changes in the outcome, BMD. BMI varies quite a lot between women, but also varies within a woman over time.

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Longitudinal data examples

- birthweight and birth order
- provides birthweights and order of infants from mothers who had 5 children in Georgia
 - interested in whether birthweight of babies changes with order
 - whether this difference depends on the *mother's age at first childbirth* or on the *weight of initial baby.*

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4000 -

3000

2000 -

1000 -

Birth Weight (g)

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Georgia Birthweights dataset





Georgia birthweight dataset

Birth order

Birth order

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Georgia Birthweights dataset

Does baseline birth weight vary by mother?
 random intercept

```
library(nlme)
```

Note: there are not enough degrees of freedom to also fit a random coefficient for birth order

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momid

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Georgia Birthweights dataset



Random effects

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Georgia Birthweights dataset

summary(gafit1)

```
## Linear mixed-effects model fit by REML
##
     Data: ga
##
          ATC
                   BIC
                          logLik
     15321 65 15341 28 -7656 826
##
##
## Random effects:
##
   Formula: ~1 | momid
           (Intercept) Residual
##
              367.2676 445.0228
## StdDev:
##
## Fixed effects: bweight ~ birthord
                  Value Std.Error DF t-value p-value
##
## (Intercept) 2995.640 41.99615 799 71.33130
## birthord
                 46.608 9.95101 799 4.68374
                                                      0
  Correlation:
##
##
            (Intr)
## birthord -0.711
##
## Standardized Within-Group Residuals:
##
           Min
                                   Med
                                                            Max
                        Q1
                                                 ۵3
## -5 26801358 -0 43683345 0 05028638 0 52703429 3 30770805
##
## Number of Observations: 1000
## Number of Groups: 200
```

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Georgia Birthweights dataset

intervals(gafit1, which = "all")

Approximate 95% confidence intervals ## Fixed effects: ## ## lower est. upper (Intercept) 2913.20418 2995.640 3078.07582 ## birthord 27.07478 46.608 66 14122 ## ## Random Effects: Level: momid ## ## lower est. upper ## sd((Intercept)) 323.1724 367.2676 417.3794 ## ## Within-group standard error: ## lower est. upper ## 423.7298 445.0228 467.3859

- Does baseline birth weight vary by mother?
 - yes: the subject variance is significantly greater than zero
 - The variance between mothers is too much to be explained by within-mother variation in birth weights

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Population Average Models

- An alternative to random / mixed-effects models that is more robust to assumptions of:
 - distribution of random effects
 - correlation structure
- Estimates correlation structure from the data rather than assuming normality
 - Requires more clusters than observations per cluster
- Estimates regression coefficients and robust standard errors
 - commonly by Generalized Estimating Equations (GEE)

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• Compare mixed model multiple linear regression:

 $E[Y_{ij}|X_{ij}] = \beta_0 + \alpha_{0j} + \beta_1 X_{ij}, \alpha_{0j} \sim N(0, \sigma)$

for subject i in group j.

• to a population average model:

$$E[Y_{ij}|X_{ij}] = \beta_0^* + \beta_1^* X_{ij}$$

- Interpretations of β^* and β are equivalent
- Numerically equivalent for linear and log-linear models (if specification of mixed model is correct), but not for logistic link.

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Fit a population average model

```
summary(gafit.gee)
 Session 10:
  Repeated
Measures and
                  ##
Longitudinal
                      GEE: GENERALIZED LINEAR MODELS FOR DEPENDENT DATA
                  ##
 Analysis II
                  ##
                      gee S-function, version 4.13 modified 98/01/27 (1998)
                  ##
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                  ## Model:
                  ## Link:
                                                 Identity
                  ## Variance to Mean Relation: Gaussian
Learning
                  ## Correlation Structure:
                                                 Exchangeable
objectives and
                  ##
outline
                  ## Call·
                  ## gee::gee(formula = bweight ~ birthord, id = momid, data = ga,
Review
                         corstr = "exchangeable")
                  ##
Non-
                  ##
hierarchical
                  ## Summarv of Residuals:
analysis
                  ##
                           Min
                                      10
                                            Median
                                                           3Q
                                                                    Max
strategies
                  ## -2795.464 -299.126
                                            48,840
                                                     341,144 1824,536
                  ##
Mixed effects
                  ##
models
                  ## Coefficients:
                  ##
                                 Estimate Naive S.E. Naive z Robust S.E. Bobust z
Longitudinal
                  ## (Intercept) 2995.640 41.973695 71.369462 38.808066 77.191170
                  ## birthord
                                   46.608 9.958128 4.680398 9.996256 4.662546
                  ##
Population
                  ## Estimated Scale Parameter: 332525.3
Average
                  ## Number of Iterations: 1
Models
                  ##
                  ## Working Correlation
                  ##
                               [.1]
                                         [.2]
                                                    [.3]
                                                              [.4]
                                                                        [.5]
                  ## [1,] 1.0000000 0.4035684 0.4035684 0.4035684 0.4035684
                  ## [2,] 0.4035684 1.0000000 0.4035684 0.4035684 0.4035684
                  ## [3,] 0.4035684 0.4035684 1.0000000 0.4035684 0.4035684
                  ## [4,] 0.4035684 0.4035684 0.4035684 1.0000000 0.4035684
                  ## [5,] 0.4035684 0.4035684 0.4035684 0.4035684 1.0000000
```

data

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Correlation assumptions for GEE

Must make some assumption about the form of correlation among grouped observations. Some options are:

- Independence:
 - no correlation between measurements within group
- Exchangeable:
 - all pairwise correlations are the same (in large-N limit)
 - nothing distinguishes one member of a cluster from another
 - appropriate in the absence of other data structures such as measurements taken through time or space
- Auto-regressive (AR-M):
 - observations taken more closely in time are more highly correlated

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Correlation assumptions for GEE (cont'd)

- Unstructured:
 - estimates a separate correlation between observations taken on each pair of "times"
- Non-stationary ("non_stat_M_dep"):
 - similar to unstructured, but assumes all correlations for pairs separated far enough in time are zero
- Stationary ("stat_M_dep"):
 - e.g. stationary of order 2: observations taken at time points 1 and 3 have the same correlation as time points 2 and 4
 - but this might be different from the correlation between observations taken at times 2 and 3
 - correlations for observations 3 or more time periods apart assumed to be zero

Fewer assumptions requires more data, and good assumptions improve results

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Help in choosing a method

Characteristic	Marginal	Fixed-effect	Mixed-effec
Distinguishes observations belonging to the same or different subjects	Yes ^a	Yes	Yes
Reliant on distribution of subject-specific effects	No	No	Yes
Subjects considered a sample from a population larger than the sample itself	Yes ^a	No	Yes
Computation handles few subjects well	No	Yes	No
Computation handles a very large number of subjects well	Yes	No	Yes
Noisy for few observations per subject	No	Yes	No
Computation handles a large number of observations per subject	Depends ^b	Yes	Yes
Accommodates variable observations per subject	Yes	Yes	Yes

Note: ^aOnly for calculation of standard errors.

^bProblems can arise under some specifications of the working covariance structure and depending on the estimation method used.

doi:10.1371/journal.pone.0146721.t002

Figure 2: Hierarchical modeling decision table from Moen et al.

Conclusions

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- Ignoring within-subject correlations can produce very wrong results, and is not always "conservative"
- Hierarchical analysis strategies are needed for any of:
 - 1 When the correlation structure is of primary interest, *e.g.* familial aggregation of disease, or consistency of treatment within centers,
 - 2 When we wish to "borrow strength" across the levels of a hierarchy in order to improve estimates, and
 - 3 When dealing with unbalanced correlated data. E.g., no requirement that each Georgia mother have exactly 5 children.
- Population average models provide a robust alternative to mixed models
 - for one level of hierarchy

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A final note on reporting results of hypothesis tests

- Include test statistic, a measure of "effect size", and test name if unclear from test statistic
- Write in plain language and let the statistics support, not lead. E.g.:
 - do: The 36 study participants had a mean age of 27.4 (SD = 12.6), significantly older than the university mean of 21.2 years (t(35) = 2.95, p = 0.01).
 - *don't*: A p-value of 0.01 indicated significant difference in age of study participants compared to all university students.
 - do: report confidence intervals where possible
- UW "Reporting Results of Common Statistical Tests in APA Format": specific examples of reporting a hypothesis test result
- STROBE guidelines for reporting observational studies: https://www.strobe-statement.org/
- A Guideline for Reporting Results of Statistical Analysis in Japanese Journal of Clinical Oncology: helpful guidelines for all parts of a manuscript

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CONGRATULATIONS!!!