Regression Discontinuity EC 607, Set 10

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Prologue

Schedule

Last time

- Introduction to selection-on-unobservables designs
- Instrumental variables (IV) and two-stage least squares (2SLS)

Today

Regression discontinuity[†]

Upcoming

Problem set 2!

† These notes largely follow notes from Michael Anderson, Imbens and Lemieux (2008), and notes from Teppei Yamamoto.

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In addition, while RD is all the rage in modern applied econometrics, Thistlewaite and Campbell wrote about it back in 1960.

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We will assume that Y_{0i} and Y_{1i} vary smoothly in X_i .

† At least in part.

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- Election runoffs are triggered if "winner" is below 50%.
- Antidiscrimination laws only apply to firms with >15 employees.
- Prisoners are eligible for early parole if some score exceeds a threshold.
- An individual is eligible for Medicare if her age is at least 65.
- You get a ticket if your speed exceeds the speed limit.
- Fifteen-percent discount at Sizzler if your age exceeds 60.
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In some cases, "treatment" is definite once we exceed the threshold.

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E.g., a politician wins an election when the difference between her vote share and her competitor's vote share exceeds zero.

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E.g., crossing some GRE threshold discontinuously increases your chances of getting into some grad schools (but doesn't guarantee admittance).

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To estimate the causal effect of D_i on Y_i , we compare the mean of Y_i just *above* the threshold to the mean of Y_i just *below* the threshold.

More formally

We can write the comparison of means **at the threshold** as

 $\lim_{x \downarrow c} E[\mathrm{Y}_i \mid \mathrm{X}_i = x] - \lim_{x \uparrow c} E[\mathrm{Y}_i \mid \mathrm{X}_i = x]$

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I.e., Because we don't observe \mathbf{Y}_{0i} for treated individuals, we extrapolate $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i = c - \varepsilon]$ to $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i = c + \varepsilon]$ for small ε .

Estimation

Thus, we estimate

$$au_{ ext{SRD}} = \lim_{x \downarrow c} E[ext{Y}_i \mid ext{X}_i = x] - \lim_{x \uparrow c} E[ext{Y}_i \mid ext{X}_i = x]$$

as the diffrence between two regression functions estimated "near" c.

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We must stay "near" to c to minimize the bias from extrapolating $E[Y_{0i} | X_i = c - \varepsilon]$ to $E[Y_{0i} | X_i = c + \varepsilon]$ (and assuming continuity).

Ex. Is there effect of a political party winning an election on that party's likelihood of winning the following election?

Is there a benefit of incumbency (at the party level)?[†]

+ Lee (2008) addresses this question via RD. Caughey and Sekhon (2011) discuss RD in this setting.
Let's start with $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$



Let's start with $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$ and $E[\mathbf{Y}_{1i} \mid \mathbf{X}_i]$.



You only win an election if your **margin of victory exceeds zero**.



 $E[\mathbf{Y}_{1i} \mid \mathbf{X}_i] - E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$ at the discontinuity gives τ_{SRD} .



Real data are a bit trickier. We must estimate $E[Y_{1i} | X_i]$ and $E[Y_{0i} | X_i]$.



Questions

1. How should we estimate $E[Y_{1i} | X_i]$ and $E[Y_{0i} | X_i]$?

2. How much data should we use—*i.e.*, what is the right **bandwidth** size?

Option 1a Linear regression with constant slopes (and all data)



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Option 1b Linear regression with constant slopes; limited to +/-50%.



Option 2a Linear regression with differing slopes (and all data)



Option 2b Linear regression with differing slopes; limited to +/- 50%.



Option 2c Linear regression with differing slopes; limited to +/- 25%.



Option 3 Differing quadratic regressions (limited to +/- 50%).



Option 4a Differing local (LOESS) regressions (limited to +/- 50%).



Option 4b Differing local (LOESS) regressions (all data).













The continuity of $E[Y_{0i} | X_i = x]$ (in x) is also very important. No sorting.



In practice

Gelman and Imbens (2018) on functional form:

We argue that controlling for global high-order polynomials in regression discontinuity analysis is a flawed approach with three major problems: it leads to noisy estimates, sensitivity to the degree of the polynomial, and poor coverage of confidence intervals. We recommend researchers instead use estimators based on local linear or quadratic polynomials or other smooth functions.

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See Imbens and Kalyanaraman (2012) for optimal bandwidth selection.

Estimation

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- 2. **Recode** X_i (the "forcing variable") as deviation from *c*, *i.e.*, $\widetilde{X}_i = X_i c$

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3. Determine a model to **estimate** $E\left[\mathbf{Y}_{i} \mid \widetilde{\mathbf{X}}_{i}\right]$ for $\widetilde{\mathbf{X}}_{i}$ above and below 0

 \circ Linear with common slopes for $E\Big[\mathrm{Y}_i \mid \widetilde{\mathrm{X}}_i < 0\Big]$ and $E\Big[\mathrm{Y}_i \mid \widetilde{\mathrm{X}}_i > 0\Big]$

- Linear/quadratic/polynomial with differing slopes
- LOESS, kernel regression, etc.

Estimation: Linear, common slope

Assumptions

- 1. $E[\mathbf{Y}_{0i} | \mathbf{X}_i = x]$ is linear in x, *i.e.*, $E[\mathbf{Y}_{0i} | \mathbf{X}_i] = \alpha + \beta \mathbf{X}_i$
- 2. Treatment effect does not depend upon \mathbf{X}_i , *i.e.*, $E[\mathbf{Y}_{1i} \mathbf{Y}_{0i} \mid \mathbf{X}_i] = au$

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Assumption $E[Y_{0i}|X_i = x]$ and $E[Y_{1i}|X_i = x]$ are linear in x, *i.e.*, $E[Y_{0i} \mid X_i] = \alpha_0 + \beta_0 X_i$ and $E[Y_{1i} \mid X_i] = \alpha_1 + \beta_1 X_i$

Now treatment effects can vary with X_i .

$$\implies E[\mathbf{Y}_{1i} - \mathbf{Y}_{0i} \mid \mathbf{X}_i] = (\alpha_1 - \alpha_0) + (\beta_1 - \beta_0) \, \mathbf{X}_i \\ E[\mathbf{Y}_i \mid \mathbf{X}_i, \mathbf{D}_i] = \mathbf{D}_i \, \underline{E}[\mathbf{Y}_{1i} \mid \mathbf{X}_i] + (1 - \mathbf{D}_i) \, \underline{E}[\mathbf{Y}_{0i} \mid \mathbf{X}_i] \\ = \alpha_0 + \beta_0 \mathbf{X}_i + (\alpha_1 - \alpha_0) \mathbf{D}_i + (\beta_1 - \beta_0) \mathbf{D}_i \mathbf{X}_i$$

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au is the LATE at $\widetilde{\mathbf{X}}_i = 0$ ($\mathbf{X}_i = c$). Estimate: Regress \mathbf{Y}_i in $\widetilde{\mathbf{X}}_i$, \mathbf{D}_i , and $\mathbf{D}_i \widetilde{\mathbf{X}}_i$.⁺

† See Appendix for omitted steps.

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As with their sharp-RD relatives, **fuzzy RDs** take advantage of a discontinuous change in treatment assignment across some threshold *c*.

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Formally,

$$0 < \lim_{x \downarrow c} \Pr(\mathrm{D}_i = 1 \mid \mathrm{X}_i = x) - \lim_{x \uparrow c} \Pr(\mathrm{D}_i = 1 \mid \mathrm{X}_i = x) < 1$$

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Ex., Exceeding a minimum GRE requirement for graduate school.

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The treatment effect defined by a fuzzy RD is the ratio of (1) to (2)

$$au_{ ext{FRD}} = rac{\lim_{x \downarrow c} E[ext{Y}_i \mid ext{X}_i = x] - \lim_{x \uparrow c} E[ext{Y}_i \mid ext{X}_i = x]}{\lim_{x \downarrow c} E[ext{D}_i \mid ext{X}_i = x] - \lim_{x \uparrow c} E[ext{D}_i \mid ext{X}_i = x]}$$

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should remind you of something–IV, where $\mathbf{Z}_i = \mathbb{I}\left\{\mathbf{X}_i \geq c\right\}$.

Accordingly, fuzzy RDs are going to have the **same requirements and interpretation as IV**.

More formally

Let $D_i(x^*)$ denote the **potential treatment status** of *i* with threshold x^* .

Why write potential treatment status D_i a function of the threshold?

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This is our monotonicity assumption for fuzzy RDs. If we raise x^* from c to $c + \epsilon$, no one joins treatment—no defiers.

Compliance

Our **compliers** in this setting are individuals such that

$$\lim_{x^{*} \downarrow \mathrm{X}_{i}} \mathrm{D}_{i}\left(x^{*}
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Thus, $au_{
m FRD}$ can be a very local LATE.

General

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If the discontinuity is not graphically apparent, most people are not going to care about the results of a few tortured regressions.

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If the discontinuity is not graphically apparent, most people are not going to care about the results of a few tortured regressions.

You're arguing you know that treatment assignment changes across the threshold. If your reader/viewer cannot see it, they're likely not going to believe your regression tables.[†]

⁺ This skepticism may be well founded. We know RDs are sensitive to functional form—and researchers have been known to *p*-hack.

Three figures

Most RD analyses will have some subset of three types of figures.

1. **Outcomes** by the running/forcing variable (X_i)

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- 3. **Density** of running/forcing variable (X_i) Is there evidence of sorting into treatment (across the threshold)?

Outcomes by running variable

These figures tend to show the average value of the outcome Y_i at evenly spaced bins of the running variable X_i .

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We have two parameter choices

1. Binwidth (h)

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that yield $K = K_0 + K_1$ bins $(k = 1, \ldots, K)$

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We then calculate summaries for each bin.

Outcomes by running variable

The bin's **number of observations**, N_k

$$N_k = \sum_{i=1}^N \mathbb{I}\left\{b_k < \mathrm{X}_i \leq b_{k+1}
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The **average treatment level** in the bin, \overline{D}_k (for fuzzy RDs)

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Outcomes by running variable

And then plot $\overline{\mathbf{Y}}_k$ against the midpoint of each bin.

Q Does crossing c clearly affect our outcome Y_i ? (Fuzzy RD reduced form)



Covariates by running variable

Now we apply the same approach to covariates (\mathbf{Z}_i) .

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Density of running variable

Finally we looking for other violations of smoothness—particularly in form gaming the threshold.

In other words: Are individuals **bunching** just above or just below the threshold?

If so, folks just below the threshold don't give us the clean counterfactual that we want for the folks just above the threshold.

McCrary (2008) suggests testing the density of X_i at c.

Density of running variable

Effectively, we can plot N_k at the midpoint of each bin.

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Density of running variable

Likely bunching (problem)

Q Is the distribution of \mathbf{X}_i smooth across c?



Additional points

- 1. No bin should cross the threshold.
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Again, if these graphs are not clear and convincing, it's going to be hard to make the case that you have a true/credible discontinuity.

Appendix

Estimation: Linear, differing slopes

Definitions of terms that magically appear

- $\widetilde{lpha}=lpha_0+eta_0 c$
- $au = (lpha_1 lpha_0) + (eta_1 eta_0) \, c$
- $ilde{eta}=(eta_1-eta_0)$

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