Regression Stuff

EC 607, Set 05

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Prologue

Schedule

Last time: Inference and simulation

Let's review using a quote from MHE

We've chosen to start with the **asymptotic approach to inference** because modern empirical work typically leans heavily on the large-sample theory that lies behind robust variance formulas. The **payoff is valid inference under weak assumptions**, in particular, a framework that makes sense for our less-than-literal approach to regression models. On the other hand, the **large-sample approach is not without its dangers**...

MHE, p. 48 (emphasis added)

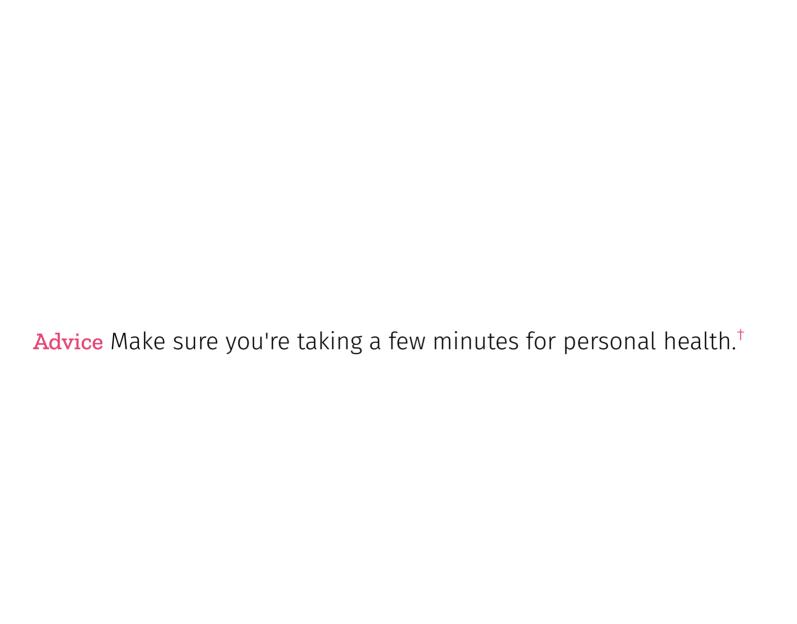
Schedule

Today

Regression and causality
Read MHE 3.2

Upcoming

Project, step 1 Assignment #2



Saturated models

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$$Wages_i = \alpha + \beta \mathbb{I}\{College\ Graduate\}_i + \varepsilon_i$$

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$$\mathrm{Wages}_i = lpha + eta_1 \, \mathbb{I}\{s_i = 1\}_i + eta_2 \, \mathbb{I}\{s_i = 2\}_i + \dots + eta_T \, \mathbb{I}\{s_i = T\}_i + arepsilon_i$$

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Here, $s_i = 0$ is our reference level; β_j is the effect of j years of schooling.

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Here, $s_i = 0$ is our reference level; β_j is the effect of j years of schooling.

$$E[\text{Wages}_i \mid s_i = j] - E[\text{Wages}_i \mid s_i = 0] = \alpha + \beta_j - \alpha = \beta_j$$

Saturated models

Q Why focus on saturated models?

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A Saturated models perfectly fit the CEF because the CEF is a linear function of the dummy variables—a special case of the linear CEF theorem.

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```
\begin{aligned} \mathbf{Wages}_i &= \alpha + \beta_1 \, \mathbb{I}\{\mathbf{College \ Graduate}\}_i + \beta_2 \, \mathbb{I}\{\mathbf{Female}\}_i \\ &+ \beta_3 \, \mathbb{I}\{\mathbf{College \ Graduate}\}_i \times \mathbb{I}\{\mathbf{Female}\}_i + \varepsilon_i \end{aligned}
```

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```

Here, the uninteracted terms $(\beta_1 \& \beta_2)$ are called **main effects**; β_3 gives the effect of the **interaction**.

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```
E[\mathrm{Wages}_i | \mathrm{College\ Graduate}_i = 0,\ \mathrm{Female}_i = 0] = \alpha
E[\mathrm{Wages}_i | \mathrm{College\ Graduate}_i = 1,\ \mathrm{Female}_i = 0] = \alpha + \beta_1
E[\mathrm{Wages}_i | \mathrm{College\ Graduate}_i = 0,\ \mathrm{Female}_i = 1] = \alpha + \beta_2
E[\mathrm{Wages}_i | \mathrm{College\ Graduate}_i = 1,\ \mathrm{Female}_i = 1] = \alpha + \beta_1 + \beta_2 + \beta_3
```

Saturated models

The CEF can take on four possible values,

```
egin{aligned} E[	ext{Wages}_i | 	ext{College Graduate}_i &= 0, 	ext{ Female}_i &= 0] &= lpha \ E[	ext{Wages}_i | 	ext{College Graduate}_i &= 1, 	ext{ Female}_i &= 0] &= lpha + eta_1 \ E[	ext{Wages}_i | 	ext{College Graduate}_i &= 0, 	ext{ Female}_i &= 1] &= lpha + eta_2 \ E[	ext{Wages}_i | 	ext{College Graduate}_i &= 1, 	ext{ Female}_i &= 1] &= lpha + eta_1 + eta_2 + eta_3 \end{aligned}
```

and the specification of our saturated regression model

$$egin{aligned} \operatorname{Wages}_i &= lpha + eta_1 \, \mathbb{I}\{\operatorname{College Graduate}\}_i + eta_2 \, \mathbb{I}\{\operatorname{Female}\}_i \ &+ eta_3 \, \mathbb{I}\{\operatorname{College Graduate}\}_i imes \mathbb{I}\{\operatorname{Female}\}_i + arepsilon_i \end{aligned}$$

does not restrict the CEF at all.

Model specification

Saturated models sit at one extreme of the model-specification spectrum, with linear, uninteracted models occupying the opposite extreme.

Saturated models

- Fit CEF (+)
- Complex (−)
 - Many dummies
 - Many interactions

Plain, linear models

- Linear approximations (–)
- Simple (+)

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Don't forget the there are many options in between—though some make less sense than others (e.g., interactions without main effects).

Model specification

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Continuous, linear probability, logged, non-negative—it works for all.

Now back to causality...

The return of causality

We've spent the last few lectures developing properties/understanding of (1) the CEF and (2) least-squares regression.

Let's return to our main goal of the course...

Q When can we actually interpret a regression as **causal**?[†]

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We've spent the last few lectures developing properties/understanding of (1) the CEF and (2) least-squares regression.

Let's return to our main goal of the course...

Q When can we actually interpret a regression as **causal**?[†]

A A regression is causal when the CEF it approximates is causal.

The return of causality

Great... thanks.

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Q So when is a CEF causal?

A First, return to the potential-outcomes framework, describing hypothetical outcomes.

A CEF is causal when it describes **differences in average potential outcomes** for a fixed reference population.

MHE, p. 52 (emphasis added)

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Great... thanks.

Q So when is a CEF causal?

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A CEF is causal when it describes **differences in average potential outcomes** for a fixed reference population.

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Let's work through this "definition" of causal CEFs with an example.

Causal CEFs

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Previously, we discussed how experiments randomly assign treatment to ensure the variable of interest is independent of potential outcomes.

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Previously, we discussed how experiments randomly assign treatment to ensure the variable of interest is independent of potential outcomes.

Now we would like to **extend this framework** to

- 1. variables that take on more than two values
- 2. situations that require us to **hold many covariates constant** in order to achieve a valid causal interpretation

Causal CEFs

The idea of holding (many) covariates constant brings us to one of the cornerstones of applied econometrics: the conditional independence assumption (CIA) (also called selection on observables).

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Definition(s)

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$$\{\mathbf{Y}_{0i}, \, \mathbf{Y}_{1i}\} \perp \!\!\! \perp \!\!\! \mathbf{D}_i | \mathbf{X}_i$$

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To see how CIA eliminates selection bias...

Selection bias
$$= E[\mathbf{Y}_{0i} \mid \mathbf{X}_i, \mathbf{D}_i = 1] - E[\mathbf{Y}_{0i} \mid \mathbf{X}_i, \mathbf{D}_i = 0]$$

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Selection bias =
$$E[\mathbf{Y}_{0i} \mid \mathbf{X}_i, \mathbf{D}_i = \mathbf{1}] - E[\mathbf{Y}_{0i} \mid \mathbf{X}_i, \mathbf{D}_i = \mathbf{0}]$$

= $E[\mathbf{Y}_{0i} \mid \mathbf{X}_i] - E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$

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 $= E[\mathbf{Y}_{0i} \mid \mathbf{X}_i] - E[\mathbf{Y}_{0i} \mid \mathbf{X}_i]$
 $= 0$

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Even randomized experiments need the CIA—e.g., the STAR experiment's within-school randomization.

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Now, Y_{si} denotes individual *i*'s outcome with *s* years of schooling.

Let each individual have her own function between schooling and earnings.

$$\mathbf{Y}_{si} \equiv f_i(\mathbf{s})$$

 $f_i(s)$ answers exactly the type of causal questions that we want to answer.

The conditional independence assumption

Extending the CIA to this multi-valued setting...

$$\mathbf{Y}_{si} \perp \!\!\! \perp s_i \mid \mathbf{X}_i$$
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If we apply the CIA to $\mathbf{Y}_{si} \equiv f_i(s)$, we define the average causal effect of a one-year increase in schooling as

$$E[f_i(\mathbf{s}) - f_i(\mathbf{s} - \mathbf{1}) \mid X_i]$$

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However, the data only contain one realization of $f_i(s)$ per i—we only see $f_i(s)$ evaluated at exactly one value of s per i, i.e., $Y_i = f_i(s_i)$.

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The CIA to the rescue! Conditional on X_i , Y_{si} and s_i are independent.

The conditional independence assumption

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$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= s] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= s-1] \ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_i, \, s_i &= s] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i, \, s_i &= s-1] \ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_i] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i] \ &= E[\mathbf{Y}_{si} - \mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i] \ &= E[f_i(s) - f_i(s-1) \mid \mathbf{X}_i] \end{aligned}$$

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The CIA to the rescue! Conditional on X_i , Y_{si} and s_i are independent.

$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, m{s}_i = m{s}] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, m{s}_i = m{s} - 1] \ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_i, \, m{s}_i = m{s}] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i, \, m{s}_i = m{s} - 1] \ &= E[\mathbf{Y}_{si} \mid \mathbf{X}_i] - E[\mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i] \ &= E[\mathbf{Y}_{si} - \mathbf{Y}_{(s-1)i} \mid \mathbf{X}_i] \ &= E[f_i(m{s}) - f_i(m{s} - 1) \mid \mathbf{X}_i] \end{aligned}$$

With the CIA, a difference in conditional averages allows causal interpretations.

The conditional independence assumption

The conditional independence assumption

$$E[\mathrm{Y}_i \mid \mathrm{X}_i,\, s_i = 12] - E[\mathrm{Y}_i \mid \mathrm{X}_i,\, s_i = 11]$$

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$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, m{s}_i = 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, m{s}_i = 11] \ &= E[f_i(12) \mid \mathbf{X}_i, \, m{s}_i = 12] - E[f_i(11) \mid \mathbf{X}_i, \, m{s}_i = 11] \end{aligned}$$

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$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 11] \ &= E[f_i(\mathbf{12}) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(\mathbf{11}) \mid \mathbf{X}_i, \, s_i &= 11] \ &= E[f_i(\mathbf{12}) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(\mathbf{11}) \mid \mathbf{X}_i, \, s_i &= 12] \end{aligned} \qquad ext{(from CIA)}$$

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$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 11] \ &= E[f_i(\mathbf{12}) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(\mathbf{11}) \mid \mathbf{X}_i, \, s_i &= 11] \ &= E[f_i(\mathbf{12}) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(\mathbf{11}) \mid \mathbf{X}_i, \, s_i &= 12] \quad & (ext{from CIA}) \ &= E[f_i(\mathbf{12}) - f_i(\mathbf{11}) \mid \mathbf{X}_i, \, s_i &= 12] \end{aligned}$$

The conditional independence assumption

$$E[Y_i \mid X_i, s_i = 12] - E[Y_i \mid X_i, s_i = 11]$$
 $= E[f_i(12) \mid X_i, s_i = 12] - E[f_i(11) \mid X_i, s_i = 11]$
 $= E[f_i(12) \mid X_i, s_i = 12] - E[f_i(11) \mid X_i, s_i = 12]$ (from CIA)
 $= E[f_i(12) - f_i(11) \mid X_i, s_i = 12]$
 $= \text{The average causal effect of graduation } for graduates$

The conditional independence assumption

$$egin{aligned} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i = 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i = 11] \ &= E[f_i(12) \mid \mathbf{X}_i, \, s_i = 12] - E[f_i(11) \mid \mathbf{X}_i, \, s_i = 11] \ &= E[f_i(12) \mid \mathbf{X}_i, \, s_i = 12] - E[f_i(11) \mid \mathbf{X}_i, \, s_i = 12] \quad & (\text{from CIA}) \ &= E[f_i(12) - f_i(11) \mid \mathbf{X}_i, \, s_i = 12] \ &= \text{The average causal effect of graduation} \ for \, graduates \ &= E[f_i(12) - f_i(11) \mid \mathbf{X}_i] \quad & (\text{CIA again}) \end{aligned}$$

The conditional independence assumption

$$\begin{split} E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 12] - E[\mathbf{Y}_i \mid \mathbf{X}_i, \, s_i &= 11] \\ &= E[f_i(12) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(11) \mid \mathbf{X}_i, \, s_i &= 11] \\ &= E[f_i(12) \mid \mathbf{X}_i, \, s_i &= 12] - E[f_i(11) \mid \mathbf{X}_i, \, s_i &= 12] \quad \text{(from CIA)} \\ &= E[f_i(12) - f_i(11) \mid \mathbf{X}_i, \, s_i &= 12] \\ &= \text{The average causal effect of graduation } for \, graduates \\ &= E[f_i(12) - f_i(11) \mid \mathbf{X}_i] \quad \text{(CIA again)} \\ &= \text{The (conditional) average causal effect of graduation } at \, X_i \end{split}$$

The conditional independence assumption

Q What about the **unconditional** average causal effect of graduation?

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A First, remember what we just showed...

The conditional independence assumption

- Q What about the **unconditional** average causal effect of graduation?
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$$E[Y_i \mid X_i, \, s_i = 12] - E[Y_i \mid X_i, \, s_i = 11] = E[f_i(12) - f_i(11) \mid X_i]$$

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$$E[Y_i \mid X_i, \, s_i = 12] - E[Y_i \mid X_i, \, s_i = 11] = E[f_i(12) - f_i(11) \mid X_i]$$

Now take the expected value of both sides and apply the LIE.

$$egin{aligned} E\Big(E[\mathbf{Y}_i\mid \mathbf{X}_i,\, oldsymbol{s}_i = \mathbf{12}] - E[\mathbf{Y}_i\mid \mathbf{X}_i,\, oldsymbol{s}_i = \mathbf{11}]\Big) \ = E\Big(E[f_i(\mathbf{12}) - f_i(\mathbf{11})\mid \mathbf{X}_i]\Big) \end{aligned}$$

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A First, remember what we just showed...

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Now take the expected value of both sides and apply the LIE.

$$egin{aligned} E\Big(E[\mathbf{Y}_i\mid\mathbf{X}_i,\,s_i=12]-E[\mathbf{Y}_i\mid\mathbf{X}_i,\,s_i=11]\Big) \ &= E\Big(E[f_i(\mathbf{12})-f_i(\mathbf{11})\mid\mathbf{X}_i]\Big) \ &= E[f_i(\mathbf{12})-f_i(\mathbf{11})] \quad ext{(Iterating expectations)} \end{aligned}$$

The conditional independence assumption

Takeaways

1. Conditional independence gives our parameters **causal interpretations** (eliminating selection bias).

The conditional independence assumption

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- 2. The interpretation changes slightly—without iterating expectations, we have **conditional average treatment effects**.
- 3. The CIA is challenging—you need to know which set of covariates (X_i) leads to as-good-as-random residual variation in your treatment.

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Takeaways

- 1. Conditional independence gives our parameters **causal interpretations** (eliminating selection bias).
- 2. The interpretation changes slightly—without iterating expectations, we have **conditional average treatment effects**.
- 3. The CIA is challenging—you need to know which set of covariates (X_i) leads to as-good-as-random residual variation in your treatment.
- 4. The idea of conditioning on observables to match *comparable* individuals introduces us to **matching estimators**—comparing groups of individuals with the same covariate values.

From the CIA to regression

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- 2. If we allow $f_i(s)$ to be nonlinear in s and heterogeneous across i, regression provides a weighted average of individual-specific differences $f_i(s) f_i(s-1)$.

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Let's start with the 'easier' case: a linear, constant-effects (causal) model.

From the CIA to regression

Let $f_i(s)$ be linear in s and equal across i except for an error term, e.g.,

$$f_i(s) = \alpha + \rho s + \eta_i \tag{A}$$

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While ρ in (A) is explicitly causal, regression-based estimates of ρ in (B) need not be causal (selection/OVB for endogenous s_i).

From the CIA to regression

Continuing with our linear, constant-effect causal model...

$$f_i(s) = \alpha + \rho s + \eta_i \tag{A}$$

Now impose the conditional independence assumption for covariates X_i .

$$\eta_i = X_i' \gamma + \nu_i \tag{C}$$

where γ is a vector of population coefficients from regressing η_i on X_i .

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where γ is a vector of population coefficients from regressing η_i on X_i .

Note Least-squares regression implies

- 1. $E[\eta_i \mid \mathbf{X}_i] = \mathbf{X}_i' \gamma$
- 2. X_i is uncorrelated with ν_i .

From the CIA to regression

$$egin{aligned} E[f_i(s) \mid \mathbf{X}_i, \, s_i] \ &= E[f_i(s) \mid \mathbf{X}_i] \end{aligned}$$
 (CIA)

From the CIA to regression

$$egin{aligned} E[f_i(oldsymbol{s}) \mid \mathbf{X}_i, \, oldsymbol{s}_i] \ &= E[f_i(oldsymbol{s}) \mid \mathbf{X}_i] \quad ext{(CIA)} \ &= E[lpha +
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ho s_i + \mathbf{X}_i' \gamma \quad \text{(Least-squares regression)} \end{aligned}$$

From the CIA to regression

Now write out the conditional expectation function of $f_i(s)$ on X_i and s_i .

$$egin{aligned} E[f_i(s) \mid \mathbf{X}_i, \, s_i] \ &= E[f_i(s) \mid \mathbf{X}_i] \quad ext{(CIA)} \ &= E[lpha +
ho s_i + \eta_i \mid \mathbf{X}_i] \ &= lpha +
ho s_i + E[\eta_i \mid \mathbf{X}_i] \ &= lpha +
ho s_i + \mathbf{X}_i' \gamma \quad ext{(Least-squares regression)} \end{aligned}$$

The CEF of $f_i(s_i)$ is linear, which means that the (right[†]) population regression will be the CEF.

 $[\]dagger$ Here, "right" means conditional on X_i .

From the CIA to regression

Thus, the linear causal (regression) model is

$$\mathbf{Y}_i = \alpha +
ho \mathbf{s}_i + \mathbf{X}_i' \gamma +
u_i$$

The residual ν_i is uncorrelated with

- 1. s_i (from the CIA)
- 2. X_i (from defining γ via the regression of η on X_i)

The coefficient ρ gives the causal effect of s_i on Y_i .

From the CIA to regression

As Angrist and Pischke note, this **conditional-independence assumption** (*a.k.a.* the selection-on-observables assumption) is the cornerstone of modern empirical work in economics—and many other disciplines.

Nearly any empirical application that wants a causal interpretation involves a (sometimes implicit) argument that **conditional on some set of covariates, treatment is as-good-as random**.

From the CIA to regression

As Angrist and Pischke note, this **conditional-independence assumption** (*a.k.a.* the selection-on-observables assumption) is the cornerstone of modern empirical work in economics—and many other disciplines.

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Part of our job: Reasoning through the validity of this assumption.

CIA example

Let's continue with the returns to graduation (G_i) .

Let's imagine

- 1. Women are more likely to graduate.
- 2. Everyone receives the same return to graduation.
- 3. Women receive lower wages across the board.

CIA example

First, we need to generate some data.

```
# Set seed
set.seed(12345)
# Set sample size
n = 1e4
# Generate data
ex_df = tibble(
  female = rep(c(0, 1), each = n/2),
  grad = runif(n, min = female/3, max = 1) %>% round(0),
  wage = 100 - 25 * female + 5 * grad + rnorm(n, sd = 3)
)
```

CIA example

Now we can estimate our naïve regression

$$\mathrm{Wage}_i = \alpha + \beta \mathrm{Grad}_i + \varepsilon_i$$

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lm(wage ~ grad, data = ex_df)

	Coef.	S.E.	t stat
Intercept	91.65	0.20	447.70
Graduate	-1.59	0.26	-6.18

CIA example

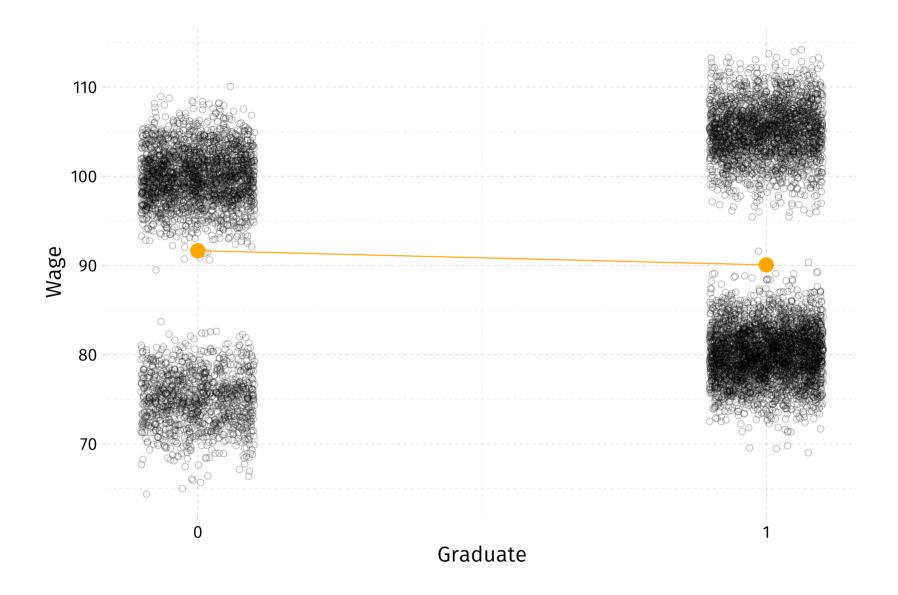
Now we can estimate our naïve regression

$$\mathrm{Wage}_i = \alpha + \beta \mathrm{Grad}_i + \varepsilon_i$$

lm(wage ~ grad, data = ex_df)

	Coef.	S.E.	t stat
Intercept	91.65	0.20	447.70
Graduate	-1.59	0.26	-6.18

Maybe we should have plotted our data...



We're still missing something...



CIA example

Now we can estimate our causal regression

$$\mathrm{Wage}_i = \alpha + \beta_1 \mathrm{Grad}_i + \beta_2 \mathrm{Female}_i + \varepsilon_i$$

CIA example

Now we can estimate our causal regression

$$\mathrm{Wage}_i = \alpha + \beta_1 \mathrm{Grad}_i + \beta_2 \mathrm{Female}_i + \varepsilon_i$$

lm(wage ~ grad + female, data = ex_df)

	Coef.	S.E.	t stat
Intercept	99.98	0.05	1868.81
Graduate	5.03	0.06	78.23
Female	-25.00	0.06	-402.64

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