## **Quantitative Macroeconomics**

# Winter 2023/24

## Week 4

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### 1. Ordinary Least Squares Estimation Of AR(p)

Consider an AR(p) model with a constant and linear term:

$$y_t = c + d \cdot t + \phi_1 y_{t-1} + \dots + \phi_p y_{t-p} + u_t = Y'_{t-1} \theta + u_t$$

where  $Y_{t-1} = (1, t, y_{t-1}, \dots, y_{t-p})$  and  $u_t \sim WN(0, \sigma_u^2)$ . The ordinary least-squares (OLS) estimator of  $\theta = (c, d, \phi_1, \dots, \phi_p)$  is

$$\hat{\theta} = \left(\sum_{t=1}^{T} Y_{t-1} Y_{t-1}'\right)^{-1} \sum_{t=1}^{T} Y_{t-1} y_t$$

Under the assumptions of stationarity and other standard regularity conditions one can derive that

$$\sqrt{T}(\hat{\theta} - \theta) \xrightarrow{d} \tilde{U} \sim N\left(0, \sigma_u^2 \ plim\left(T^{-1}\sum_{t=1}^T Y_{t-1}Y_{t-1}'\right)^{-1}\right)$$

The residual variance may be estimated consistently by

$$\hat{\sigma}_u^2 = \frac{1}{T - length(\theta)} \sum_{t=1}^T \hat{u}_t^2$$

where  $\hat{u}_t = y_t - Y'_{t-1}\hat{\theta}$  are the OLS residuals.

- 1. Write a function OLS = ARpOLS  $(y, p, const, \alpha)$  that takes as inputs a data vector y and number of lags p. The input const is 1 if there is a constant in the model, 2 if there is a constant and a linear trend. The function outputs a structure OLS, which contains the OLS estimates of  $\theta$ , its standard errors, t-statistics and p-values given significance value  $\alpha$ , as well as the OLS estimate of  $\sigma_u$ .
- 2. Load simulated data for an AR(4) model given in the CSV file AR4.csv. Estimate an AR(4) model with a constant term using your ARpOLS function.

#### Readings

• Lütkepohl (2004)

#### 2. Maximum Likelihood Estimation Of Gaussian AR(p)

Consider an AR(p) model with a constant and linear trend:

$$y_t = c + d \cdot t + \phi_1 y_{t-1} + \dots + \phi_p y_{t-p} + u_t = Y_{t-1}\theta + u_t$$

where  $Y_{t-1} = (1, t, y_{t-1}, \ldots, y_{t-p})$  is the matrix of regressors,  $\theta = (c, d, \phi_1, \ldots, \phi_p)$  the parameter vector and the error terms  $u_t$  are white noise and normally distributed, i.e.  $u_t \sim N(0, \sigma_u^2)$  and  $E[u_t u_s] = 0$ for  $t \neq s$ . If the sample distribution is known to have probability density function  $f(y_1, \ldots, y_T)$ , an estimation with Maximum Likelihood (ML) is possible. To this end, we decompose the joint distribution by

 $f(y_1,\ldots,y_T|\theta,\sigma_u^2) = f_1(y_1|\theta,\sigma_u^2) \times f_2(y_2|y_1,\theta,\sigma_u^2) \times \cdots \times f_T(y_T|y_{T-1},\ldots,y_1,\theta,\sigma_u^2)$ 

Then the log-likelihood is

$$\log f(y_1,\ldots,y_T|\theta,\sigma_u^2) = \sum_{t=1}^T \log f_t(y_t|y_{t-1},\ldots,y_1,\theta,\sigma_u^2)$$

Let's denote the values that maximize the log-likelihood as  $\tilde{\theta}$  and  $\tilde{\sigma}_u^2$ . ML estimators have (under general assumptions) an asymptotic normal distribution

$$\sqrt{T} \begin{pmatrix} \tilde{\theta} - \theta \\ \tilde{\sigma}_u^2 - \sigma_u^2 \end{pmatrix} \stackrel{d}{\to} U \sim N(0, I_a(\theta, \sigma_u^2)^{-1})$$

where  $I_a(\theta, \sigma_u)$  is the asymptotic information matrix. Recall that the asymptotic information matrix is the limit of minus the expectation of the Hessian of the log-likelihood divided by the sample size.

$$I_a(\theta, \sigma_u^2) = \lim_{T \to \infty} -\frac{1}{T} E \begin{pmatrix} \frac{\partial^2 \log l}{\partial \theta^2} & \frac{\partial^2 \log l}{\partial \theta \partial \sigma_u^2} \\ \frac{\partial^2 \log l}{\partial \sigma_u^2 \partial \theta} & \frac{\partial^2 \log l}{\partial (\sigma_u^2)^2} \end{pmatrix}$$

- 1. First consider the case of p = 1
  - a) Derive the exact log-likelihood function for the AR(1) model with  $|\theta| < 1$  and d = 0:

$$y_t = c + \theta y_{t-1} + u_t$$

- b) Why do we often look at the log-likelihood function instead of the actual likelihood function?
- c) Regard the value of the first observation as deterministic or, equivalently, note that its contribution to the log-likelihood disappears asymptotically. Maximize analytically the conditional log-likelihood to get the ML estimators for  $\theta$  and  $\sigma_u$ . Compare these to the OLS estimators.
- 2. Now consider the general AR(p) model.
  - a) Write a function logLikeARpNorm(x, y, p, const) that computes the value of the log-likelihood conditional on the first p observations of a Gaussian AR(p) model, i.e.

$$\log l(\theta, \sigma_u) = -\frac{T-p}{2}\log(2\pi) - \frac{T-p}{2}\log(\sigma_u^2) - \sum_{t=p+1}^T \frac{u_t^2}{2\sigma_u^2}$$

where  $x = (\theta', \sigma_u)'$ , y denotes the data vector, p the number of lags and *const* is equal to 1 if there is a constant, and equal to 2 if there is a constant and linear trend in the model.

b) Write a function ML = ARpML( $y, p, const, \alpha$ ) that takes as inputs a data vector y, number of lags p and const = 1 if the model has a constant term or const = 2 if the model has a constant term and linear trend.  $\alpha$  denotes the significance level. The function computes

- the maximum likelihood estimates of an AR(p) model by numerically minimizing the negative conditional log-likelihood function using e.g. fminunc
- the standard errors by means of the asymptotic covariance matrix, i.e. the inverse of the hessian of the negative log-likelihood function (hint: gradient-based optimizers also output the hessian)

Save all results into a structure "ML" containing the estimates of  $\theta$ , its standard errors, t-statistics and p-values as well as the ML estimate of  $\sigma_u$ .

c) Load simulated data given in the CSV file AR4.csv and estimate an AR(4) model with constant term. Compare your results with the OLS estimators from the previous exercise.

#### Readings

• Lütkepohl (2004).

## 3. Maximum Likelihood Estimation Of Laplace AR(p)

Consider the AR(1) model with constant

$$y_t = c + \phi y_{t-1} + u_t$$

Assume that the error terms  $u_t$  are i.i.d. Laplace distributed with known density

$$f_{u_t}(u) = \frac{1}{2} \exp\left(-|u|\right)$$

Note that for the above parametrization of the Laplace distribution we have that  $E(u_t) = 0$  and  $Var(u_t) = 2$ , so we are only interested in estimating c and  $\phi$  and not the standard deviation of  $u_t$  as it is fixed.

- 1. Derive the log-likelihood function conditional on the first observation.
- 2. Write a function that calculates the conditional log-likelihood of c and  $\phi$ .
- 3. Load the dataset given in the CSV file LaPlace.csv. Numerically find the maximum likelihood estimates of c and  $\phi$  by minimizing the negative conditional log-likelihood function.
- 4. Compare your results with the maximum likelihood estimate under the assumption of Gaussianity. That is, redo the estimation by minimizing the negative Gaussian log-likelihood function.

#### Readings

• Lütkepohl (2004)

## References

Lütkepohl, Helmut (2004). "Univariate Time Series Analysis". In: Applied Time Series Econometrics. Ed. by Helmut Lütkepohl and Markus Krätzig. 1st ed. Cambridge University Press, pp. 8–85. ISBN: 978-0-521-83919-8 978-0-521-54787-1 978-0-511-60688-5. URL: https://doi.org/10.1017/ CB09780511606885.003.

### A. Solutions

#### 1 Solution to Ordinary Least Squares Estimation Of AR(p)

```
progs/matlab/ARpOLS.m
 1.
   function OLS = ARpOLS(y,p,const,alph)
 1
 2
   % OLS = ARpOLS(y,p,const,alph)
 3
   % —
   % OLS regression of AR(p) model:
 4
   % y_t = c + d*t + theta_1*y_{t-1} + ... + theta_p*y_{t-p} + u_t
 5
 6
   % with white noise u_t \sim (0, sigma_u)
 7
   % ____
   % INPUT
8
   % — y
                [Tx1] data vector of dimension T
9
                [scalar] number of lags
10
   %
      — p
11
   % — const [scalar] 1 constant; 2 constant and linear trend in model
12
   %
          - alpha [scalar] significance level for t statistic and p value
13
   % —
14
   % OUTPUT
15
   %

    OLS: structure including estimation results

16 %
         — T_eff
                        [scalar] effective sample size used in estimation
                        [(const+p)x1] estimate of coefficients
17 %
         — thetahat
        - sd_thetahat [(const+p)x1] estimate of standard error of coefficients
18 %
                        [(const+p)x1] t statistics
19 %
        — tstat
20
                       [(const+p)x1] p values of H_0: thetahat = 0
   %
        — pvalues
                       [scalar] estimate of standard deviation of error term
21 %
        — siguhat
       u
22
   %
       — theta_ci [(const+p)x2] (1—alph)% confidence intervall for theta
       given significance level alph
23
      — resid
                 [T_eff x 1] residuals
   %
24 %-
25
   % Calls
26 \ | \$ - lagmatrix.m (Econometrics Toolbox)
27
   % —
   % Willi Mutschler, November 9, 2022
28
29
   % willi@mutschler.eu
30 %-
32 | T = size(y, 1);
                            % sample size
33 \mid T_eff = T_p;
                            % effective sample size used in estimation
34 Y = lagmatrix(y,1:p);
                            % create matrix with lagged variables
35 if const==1
                            % add constant term
36
      Y = [ones(T,1) Y];
                            % add constant term and time trend
37
   elseif const==2
38
       Y = [ones(T,1) transpose(1:T) Y];
39 end
40 Y = Y((p+1):end,:);
                            % get rid of initial p observations
                            % get rid of initial p observations
41 y = y(p+1:end);
42
43 YtYinv = inv(Y'*Y);
44 | thetahat = YtYinv*(Y'*y); % OLS estimator of coefficients
45 | yhat = Y*thetahat; % predicted values
                           % residuals
46 | uhat = y - yhat;
```

```
47 utu = uhat'*uhat; % sum of squared residuals
48
49 var_uhat = utu/(T_eff_p_const);
                                        % variance of error term
50 siguhat = sqrt(var_uhat);
                                        % standard deviaiont of error term
51 var_thetahat = diag(var_uhat*(YtYinv)); % variance of coefficients
52 |sd_thetahat = sqrt(var_thetahat); % standard error of coefficients
53
54 tstat = thetahat./sd_thetahat;
                                        % t—statistics
55 tcrit = -tinv(alph/2,T_eff-p-const);
                                        % critical value
56 pvalues = tpdf(tstat,T_eff-p-const);
                                        % p—value
57 \% confidence interval
58 theta_ci=[thetahat-tcrit.*sd_thetahat, thetahat+tcrit.*sd_thetahat];
59
60 % Store into output structure
61 \quad \text{OLS.T_eff} = \text{T_eff};
62 OLS.thetahat
                  = thetahat;
63 OLS.siguhat
                  = siguhat;
64 OLS.sd_thetahat = sd_thetahat;
65 OLS.tstat
                   = tstat;
66 OLS.pvalues
                  = pvalues;
67 OLS.theta_ci
                  = theta_ci;
68 OLS.resid
                   = uhat;
69 end
```

#### progs/matlab/AR4OLS.m

```
1
   % —
   % Estimation of AR(4) model with constant using OLS on simulated data
2
   % ____
3
4 % Willi Mutschler, November 9, 2022
5 % willi@mutschler.eu
6 %-
7
8 % Housekeeping
9 clearvars; clc; close all;
10 % load data
11 AR4 = importdata("../../data/AR4.csv");
12 y = AR4.data;
                                       % vector with data
13 p = 4;
                                       % set number of lags
14 const = 1;
                                       % model with constant
15 alph = 0.05;
                                      % significance level
16 |OLS = ARpOLS(y,p,const,alph); % estimate model using ARpOLS function
17
18 % Display results and compare to true values
19 |TrueVals = [1; 0.51; -0.1; 0.06; -0.22; 0.5];
20 result = table([OLS.thetahat;OLS.siguhat],TrueVals);
21 | result.Properties.VariableNames = { 'OLS_Estimate', 'True_Values' };
22
   result.Properties.RowNames = {'c', '\phi_1', '\phi_2', '\phi_3', '\phi_4', '\sigma_u'
       };
23 disp(result)
```

#### 2. 2 Solution to Maximum Likelihood Estimation Of Gaussian AR(p)

- 1. Let's first consider the AR(1) model
  - a) The first observation  $y_1$  is a random variable with mean and variance equal to:

$$E[y_1] = \mu = \frac{c}{1-\theta}$$
 and  $Var[y_1] = \frac{\sigma_u^2}{1-\theta^2}$ 

Since the errors are Gaussian,  $y_1$  is also Gaussian, i.e.  $y_1 \sim N\left(\frac{c}{1-\theta}, \frac{\sigma_u^2}{1-\theta^2}\right)$ . The pdf is thus:

$$f_1(y_1|\theta, \sigma_u^2) = \frac{1}{\sqrt{2\pi}\sqrt{\sigma_u^2/(1-\theta^2)}} \exp\left\{-\frac{1}{2} \frac{[y_1 - (c/(1-\theta))]^2}{\sigma_u^2/(1-\theta^2)}\right\}$$

The second observation  $y_2$  conditional on  $y_1$  is given by  $y_2 = c + \theta y_1 + u_2$ . Conditional on  $y_1, y_2$  is thus the sum of a deterministic term  $(c + \theta y_1)$  and the  $N(0, \sigma_u^2)$  variable  $u_2$ . Hence:

$$y_2|y_1 \sim N(c + \theta y_1, \sigma_u^2)$$

and the pdf is given by:

$$f_2(y_2|y_1, \theta, \sigma_u^2) = \frac{1}{\sqrt{2\pi\sigma_u^2}} \exp\left\{-\frac{1}{2} \frac{[y_2 - c - \theta y_1]^2}{\sigma_u^2}\right\}$$

The joint density of observations 1 and 2 is then just:

$$f(y_2, y_1 | \theta, \sigma_u^2) = f_2(y_2 | y_1, \theta, \sigma_u^2) \cdot f_1(y_1 | \theta, \sigma_u^2)$$

In general the value of  $y_1, y_2, ..., y_{t-1}$  matter for  $y_t$  only through the value  $y_{t-1}$  and the density of observation t conditional on the preceding t-1 observations is given by

$$f_t(y_t|y_{t-1}, \theta, \sigma_u^2) = \frac{1}{\sqrt{2\pi\sigma_u^2}} \exp\left\{-\frac{1}{2} \frac{[y_t - c - \theta y_{t-1}]^2}{\sigma_u^2}\right\}$$

The likelihood of the complete sample can thus be calculated as:

$$f(y_T, y_{T-1}, ..., y_1 | \theta, \sigma_u^2) = f_1(y_1 | \theta, \sigma_u^2) \cdot \prod_{t=2}^T f_t(y_t | y_{t-1}, \theta, \sigma_u^2)$$

The log-likelihood is therefore

$$\log l(\theta, \sigma_u^2) = \log f_1(y_1|\theta, \sigma_u^2) + \sum_{t=2}^T \log f_t(y_t|y_{t-1}, \theta, \sigma_u^2)$$
  
=  $-\frac{1}{2} \log(2\pi) - \frac{1}{2} \log(\sigma_u^2/(1-\theta^2)) - \frac{(y_1 - (c/(1-\theta))^2)}{2\sigma_u^2/(1-\theta^2)}$   
-  $((T-1)/2) \log(2\pi) - ((T-1)/2) \log(\sigma_u^2) - \sum_{t=2}^T \frac{(y_t - c - \theta y_{t-1})^2}{2\sigma_u^2}$ 

b) Theoretically it does not matter whether we consider the log-likelihood or the actual likelihood function, as the value that maximizes the likelihood also maximize the log-likelihood, because the log is a monotone transformation. However, it is usually easier to work with sums instead of products theoretically, e.g. the LLN or CT are based on sums. Computationally, working with products is typically impossible as the resulting values very quickly surpass machine precision (they quickly go to  $\pm \infty$ ); working with sums does not have this problem. So from a computational perspective we will exclusively work with the log-likelihood function.

c) Discarding the first observation, the conditional log-likelihood is given by:

$$\log l^{c}(\theta, \sigma_{u}^{2}) = -((T-1)/2)\log(2\pi) - ((T-1)/2)\log(\sigma_{u}^{2}) - \sum_{t=2}^{T} \frac{(y_{t} - c - \theta y_{t-1})^{2}}{2\sigma_{u}^{2}}$$
$$= -((T-1)/2)\log(2\pi) - ((T-1)/2)\log(\sigma_{u}^{2}) - \sum_{t=2}^{T} \frac{u_{t}^{2}}{2\sigma_{u}^{2}}$$

Note that the first two sums do not depend on  $\theta$ ; thus, when maximizing  $\log l^c(\theta, \sigma_u^2)$  with respect to  $\theta$ , we are basically minimizing the squared residuals, which will simply yield the OLS estimator. The estimator for the variance, however, is different, as we are dividing the sum of sqared residuals  $(\sum_{t=2}^{T} u_t^2)$  by  $T^{eff} = (T-1)$  when doing ML instead of by  $T^{eff} - 1$  when doing OLS. Obviously, for large T this does not matter.

```
progs/matlab/logLikeARpNorm.m
```

```
function loglik=logLikeARpNorm(x,y,p,const)
1
2
   % loglik=logLikeARpNorm(x,y,p,const)
3
   % —
   % Computes the conditional log likelihood function of Gaussian AR(p) model:
4
   % y_t = c + d*t + theta_1*y_{t-1} + ... + theta_p*y_{t-p} + u_t
5
   % with u_t \sim N(0, sig_u)
6
7
   % –
   % INPUT
8
9
   %
       — x
                 [(const+p+1)x1] vector of coefficients, i.e. [c,d,theta_1,...,
       theta_p,sig_u]'
                                   data vector of dimension T
10
   %
       — y
                 [Tx1]
11
   %
       — p
                 [scalar]
                                   number of lags
12
   %
       — const [scalar]
                                   1 constant; 2 constant and linear trend in model
13
   % –
14
   % OUTPUT
   %
15
           - loglik [double]
                                       value of Gaussian log—likelihood
16
   % _
17
   % Calls
18
   % — lagmatrix.m (requires Econometrics Toolbox)
19
   % —
20
   % Willi Mutschler, November 14, 2023
21
   % willi@mutschler.eu
22
   % ____
   penalizedLikelihood = -1e10; % very small number to penalize likelihood, e.g. -
23
       Inf
24
25
   theta = x(1:(const+p)); % [c d theta']
26
   sig_u = x(const+p+1); % standard deviation of error term sig_u
   % make sure sig_u is positive
27
28
   if sig_u <= 0</pre>
29
        loglik = penalizedLikelihood;
30
                                % this ends the current function call
        return
31
   else
       T = size(y,1);
                                 % sample size
33
       Y = lagmatrix(y, 1:p);
                                % create matrix with lagged variables
34
       if const == 1
                                 % add constant
35
            Y = [ones(T,1) Y];
36
        elseif const == 2
                                % add constant and time trend
```

```
37
           Y = [ones(T,1) transpose(1:T) Y];
38
       end
39
       Y = Y((p+1):end,:);
                               % get rid of initial observations
40
       y = y(p+1:end);
                               % get rid of initial observations
41
42
       uhat = y - Y * theta;
                               % ML residuals
43
       utu = uhat'*uhat;
                               % ML sum of residuals
44
45
       % compute the conditional log likelihood
46
       loglik = -log(2*pi)*(T-p)/2 - log(sig_u^2)*(T-p)/2 - utu/(2*sig_u^2);
47
   end
48
49
   % if anything goes wrong set value very small
   if isnan(loglik) || isinf(loglik) || ~isreal(loglik)
50
       loglik = penalizedLikelihood;
51
52 end
53
54 end % function end
```

progs/matlab/ARpML.m

```
3.
   function ML = ARpML(y,p,const,alph)
 1
2
   % ML = ARpML(y,p,const,alph)
3 % ----
   % Maximum Likelihood Estimation of Gaussian AR(p) model:
4
   % y_t = c + d*t + theta_1*y_{t-1} + ... + theta_p*y_{t-p} + u_t
5
   % with u_t ~ iid N(0,sig_u)
6
7
   % _____
   % INPUTS
8
   % – y [Tx1] dependent variable vector
9
10
   % — p [scalar] number of lags
   %
      — const [scalar] 0 no constant; 1 constant; 2 constant and linear trend
11
12
   % — alph [scalar] significance level for t statistic
   % —
13
14 % OUTPUT
15 %
         - ML: structure including estimation results
16
   %
        — T_eff
                          [scalar] effective sample size used in estimation
17 %
        — thetatilde
                         [(const+p)x1] estimate of coefficients
18
   %
         - sd_thetatilde [(const+p)x1] estimate of standard error of coefficients
   %
19
                         [(const+p)x1] t statistics given alph as significance
        — tstat
      level
                         [(const+p)x1] p values of H_0: thetahat = 0
20
       — pvalues
   %
21 %
        — sigutilde
                         [scalar]
                                        estimate of standard deviation of error
      term u
22
                                        estimate of standard error of standard
   %
       — sd_sigutilde
                        [scalar]
      deviation of error term u
   %
23
                         [(const+p)x2] (1—alph)% confidence intervall for theta
       — theta_ci
      given significance level alph
24
   %
       — logl
                          [double]
                                        value of maximized log likelihood
25
   % _
26 % CALLS
27 8 — logLikeARpNorm.m : Computes log likelihood function of Gaussian AR(p)
```

```
28 %
        — hessian_numerical.m : Computes second order partial derivatives using
       numerical differentiation
29
   % –
30 % Willi Mutschler, November 9, 2022
31
   % willi@mutschler.eu
32
   % ___
33
34 | T = size(y, 1);
                            % sample size
35
36
   % Optimization with fminunc which finds the minimum of negative log-likelihood
37 | f = @(x) -1*logLikeARpNorm(x,y,p,const); % use function handle to hand over
       additional parameters and multiply by -1 for negative log-likelihood
38 % start values
   x0 = randn(p+const+1,1); % randomize
39
40 |x0(end) = abs(x0(end)); % make sure initial sig_u is positive
41
42 [x,fval,exitflag,output,grad,hess] = fminunc(f,x0);
43
   % alternatively use hessian_numerical.m that does two—sided finite difference
       computation of hessian
44
   % [x,fval] = fminunc(f,x0);
45 |hess = reshape(hessian_numerical(f,x),length(x),length(x)); %alternatively use
       output argument of fminunc
46
47
   thetatilde = x(1:p+const);
                                       % estimated coefficient values
48 sigutilde = x(end);
                                       % estimated standard daviation of error term
                                       % estimated covariance matrix of coefficients
49 V = inv(hess);
        and (log of) standard deviation of error
                                       % estimated standard error vector
50 | sd = sqrt(diag(V));
   sd_thetatilde = sd(1:p+const);
                                       % estimated standard errors of coefficients
51
52 sd_sigutilde = sd(end);
                                       % estimated standard error of standard
       deviation of error term
53 T_eff = T_p;
                                       % effective sample size used in estimation
54 |\log| = -fval;
                                       % value of maximized log likelihood
55
   tstat = thetatilde./sd_thetatilde; % t—statistic
56 tcrit = -tinv(alph/2,T_eff-p);
                                       % critical value from t—distribution
57
   pvalues = tpdf(tstat,T_eff-p);
                                       % p—values from t—distribution
58
59
   % confidence interval for coefficients given significance level alph
60 | theta_ci=[thetatilde-tcrit.*sd_thetatilde, thetatilde+tcrit.*sd_thetatilde];
61
62
   % Store into output structure
63 ML.T_eff
                     = T_eff;
64
   ML.logl
                     = logl;
65 ML.thetatilde
                    = thetatilde;
66 ML.sigutilde
                    = sigutilde;
67 ML.sd_thetatilde = sd_thetatilde;
68 ML.sd_sigutilde = sd_thetatilde;
69
   ML.tstat
                     = tstat;
70 ML.pvalues
                     = pvalues;
71 ML.theta_ci
                     = theta_ci;
72
73 end %function end
```

progs/matlab/AR4ML.m

```
1
   %
2
   % Estimation of AR(4) model with constant using ML on simulated data
3
   % _
   % Willi Mutschler, November 9, 2022
4
   % willi@mutschler.eu
5
6
   % _
7
8 % Housekeeping
9 clearvars; clc; close all;
10 % load data
11 AR4 = importdata("../../data/AR4.csv");
12
   y = AR4.data;
13
   p = 4;
                                       % set number of lags
14 | const = 1;
                                       % model with constant
                                       % significance level
15
   alph = 0.05;
16 ML = ARpML(y,p,const,alph);
                                       % estimate model using ARpML function
17
18 % Display results and compare to true values
19 |TrueVals = [1; 0.51; -0.1; 0.06; -0.22; 0.5];
   result = table([ML.thetatilde;ML.sigutilde],TrueVals);
20
21
   result.Properties.VariableNames = {'ML_Estimate', 'True_Values'};
22
   result.Properties.RowNames = {'c', '\phi_1', '\phi_2', '\phi_3', '\phi_4', '\sigma_u'
       };
23 disp(result)
24
25 % Compare to OLS estimates
26
   OLS = ARpOLS(y,p,const,alph);
27 disp([TrueVals [OLS.thetahat; OLS.siguhat] [ML.thetatilde; ML.sigutilde]]);
```

The estimates for the coefficients are the same, but slightly different for the standard deviation of the error term.

#### 4. 3 Solution to Maximum Likelihood Estimation Of Laplace AR(p)

- 1. Computation of the conditional expectation and variance:
  - $E[y_t|y_{t-1}] = c + \phi y_{t-1}$
  - $Var[y_t|y_{t-1}] = var(u_t) = 2$

Hence the conditional density is

$$f_t(y_t|y_{t-1}; c, \phi) = \frac{1}{2} \cdot e^{-|y_t - (c + \phi y_{t-1})|} = \frac{1}{2} \cdot e^{-|u_t|}$$

The conditional log-likelihood function is therefore given by

$$\log L(y_2, \dots, y_T; c, \phi) = -(T-1) \cdot \log(2) - \sum_{t=2}^T |u_t|$$

```
progs/matlab/logLikeARpLaplace.m
 2.
   function loglik=logLikeARpLaplace(x,y,p,const)
 1
 2
   % loglik=logLikeARpLaplace(x,y,p,const)
 3
   % ____
   % Computes the conditional log likelihood function of Laplace AR(p) model:
 4
   % y_t = c + d*t + theta_1*y_{t-1} + ... + theta_p*y_{t-p} + u_t
 5
   % with u_t \sim \text{Laplace distributed with } E(u_t)=0, Var(u_t)=2 (known variance)
 6
 7
   % —
 8
   % INPUT
   % — X
 9
                [(const+p)x1] vector of coefficients, i.e. [c,d,theta_1,...,theta_p
       11
10
   %
      — y
                [Tx1]
                               data vector of dimension T
11
   %
                [scalar]
                               number of lags
      — p
12
   % — const [scalar]
                               1 constant; 2 constant and linear trend in model
   % ___
13
14 % OUTPUT
15 %
         — loglik [double]
                                  value of Laplace log—likelihood
16
   % —
17
   % Calls
18
   % — lagmatrix.m (requires Econometrics Toolbox)
19
   % ___
20
   % Willi Mutschler, November 9, 2022
21
   % willi@mutschler.eu
22
   %____
23
24 | theta = x;
                           % AR coefficients
25 | T = size(y,1);
                          % sample size
26
27 |Y = lagmatrix(y,1:p); % create matrix with lagged variables
28
   if const == 1
                           % add constant
29
       Y = [ones(T,1) Y];
30 elseif const == 2
                           % add constant and time trend
31
       Y = [ones(T,1) transpose(1:T) Y];
32
   end
33 Y = Y((p+1):end,:); % get rid of initial observations
34 | y = y(p+1:end);
                         % get rid of initial observations
```

```
36
   uhat = y - Y * theta;
                          % ML residuals
37
38
   % compute the conditional log likelihood
39
   loglik = -log(2)*(T-p) - sum(abs(uhat));
40
   if isnan(loglik) || isinf(loglik) || ~isreal(loglik)
41
42
       loglik = -le10; % if anything goes wrong set value very small, can also
           use -Inf
43
   end
44
45 \mid end % function end
```

progs/matlab/ARpMLLaPlace.m

```
function ML = ARpMLLaplace(y,p,const,alph)
1
2
   % ML = ARpMLLaplace(y,p,const,alph)
3
   % ____
   % Maximum Likelihood estimation of Laplace AR(p) model:
4
   % y_t = c + d*t + theta_1*y_{t-1} + ... + theta_p*y_{t-p} + u_t
5
6
   % with u_t \sim \text{Laplace distributed with } E(u_t)=0, Var(u_t)=2
7
   % —
   % INPUTS
8
9
   % — y
                [Tx1]
                          dependent variable vector
                [scalar] number of lags
10
   % — p
11
   %
      - const [scalar] 0 no constant; 1 constant; 2 constant and linear trend
12
   % — alph
                [scalar] significance level for t statistic
13
   % _
   % OUTPUT
14
   %
15
          - ML: structure including estimation results
16
   %
                                          effective sample size used in estimation
         — T_eff
                          [scalar]
   %
                          [(const+p)x1] estimate of coefficients
17
         — thetatilde
         - sd_thetatilde [(const+p)x1] estimate of standard error of coefficients
18
   %
19
   %
        — tstat
                          [(const+p)x1] t statistics given alph as significance
       level
20
                           [(const+p)x1] p values of H_0: thetahat = 0
   %
        — pvalues
   %
21
                           [(const+p)x2] (1—alph)% confidence intervall for theta
         — theta_ci
       given significance level alph
22 %
         — loql
                           [double]
                                          value of maximized log likelihood
23
   % -
24 % CALLS
25 %
        — logLikeARpLaPlace : Computes log likelihood function of Laplace AR(p)
26 %
        — hessian_numerical.m : Computes second order partial derivatives using
       numerical differentiation
27
28
   % Willi Mutschler, November 9, 2022
29
   % willi@mutschler.eu
31
32 | T = size(y, 1);
                                       % sample size
33 \times 0 = randn(p+const,1);
                                       \% randomize start values, note that sig_u is
       known
34 |%x0 = [1;0.8]; %true values
35 \mid% Optimization with fminunc which finds the minimum of negative log—likelihood
```

```
36
   fun = @(x)-1*\logLikeARpLaplace(x,y,p,const); % use function handle to hand over
37
                                              % additional parameters to negative
38
                                              % of logLikeARpLaplace
39 | \$[x, fval, exitflag, output, grad, hessian] = fminunc(fun, x0); \$ the hessian might be
       badly shaped, better use hessian_numerical.m
40 | [x, fval] = fminunc(fun, x0);
41
   hess = reshape(hessian_numerical(fun,x),length(x),length(x));
42
43 thetatilde = x;
                                       % estimated coefficient values
                                       % estimated covariance matrix of coefficients
44 V = inv(hess);
                                       % estimated standard error of coefficients
45 sd_thetatilde = sqrt(diag(V));
46 | T_eff = T_p;
                                       % effective sample size used in estimation
                                       % value of maximized log likelihood
47 |logl = -fval;
48 tstat = thetatilde./sd_thetatilde; % t—statistic
49 |tcrit = -tinv(alph/2,T_eff-p);
                                       % critical value from t-distribution
                                       % p-values from t-distribution
50 pvalues = tpdf(tstat,T_eff—p);
51
52
   % confidence interval given signficance level alph
53
   theta_ci=[thetatilde_tcrit.*sd_thetatilde, thetatilde+tcrit.*sd_thetatilde];
54
55 % Store into output structure
56 ML.T_eff
                   = T_eff;
57 ML.logl
                   = logl;
58 ML.thetatilde = thetatilde;
59 ML.sd_thetatilde = sd_thetatilde;
60 ML.tstat
                   = tstat;
                   = pvalues;
61 ML.pvalues
62 ML.theta_ci
                  = theta_ci;
63
64 end %function end
```

#### progs/matlab/AR1MLLaPlace.m

```
1
   % __
2 \% Estimate Laplace AR(1) model with constant and known variance of errors
   % with Maximum Likelihood on simulated data
3
4 % ---
   % Willi Mutschler, November 9, 2022
5
6 % willi@mutschler.eu
7
   %____
8
9 clearvars; clc; close all;
                                               % housekeeping
10 y = importdata('../../data/LaPlace.csv'); % load data
11 y = y.data;
                                               % focus on numerical values
12
   p = 1;
                                               % set number of lags
13 const = 1;
                                               % model with constant
14 alph = 0.05;
                                               % significance level
15 MLLaPlace = ARpMLLaplace(y,p,const,alph); % estimate model using ARpMLLaplace
16
17 \ Display results and compare to true values
18 |TrueVals = [1; 0.8];
19 | result = table(TrueVals,MLLaPlace.thetatilde,MLLaPlace.sd_thetatilde);
20 | result.Properties.VariableNames = { 'True_Values', 'ML_Estimate', 'STD_ERR'};
```

```
21 result.Properties.RowNames = {'c','\phi'};
22 disp(result)
23 
24 % Compare to Gaussian ML estimates
25 MLGaussian = ARpML(y,p,const,alph); % estimate model using ARpML
26 disp([TrueVals MLLaPlace.thetatilde MLGaussian.thetatilde]);
27 disp([MLLaPlace.sd_thetatilde MLGaussian.sd_thetatilde]);
```

**3.** Note that the values are very close to each other. Maximizing the Gaussian likelihood, even though the underlying distribution is not Gaussian, is also known as pseudo-maximum likelihood or quasi-maximum likelihood. It usually performs surprisingly well if you cannot pin down the underlying distribution.