

Solutions to problems from: A Crash Course in Practical Data Analysis*

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*For the latest version, visit <https://github.com/sashahafner/CCPDA>

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1 Packages and functions

```
source('functions/dfsumm.R')
```

```
library(tidyr)
library(dplyr)
library(ggplot2)
```

2 Problem 1. Inoculum effects on BMP

Koch et al. [2017] studied the effect of inoculum origin on biochemical methane potential (BMP) for four substrates. Data are given in the file BMP_inoc.csv, where the unit of observation is a single BMP bottle. Take a look at the data and answer these questions:

1. Did BMP depend on inoculum type?
2. Did any effect vary by substrate?

The original data are in a intermediate structure, with replicates across columns.

```
bi <- read.csv('data/BMP_inoc.csv')
```

```
bi
```

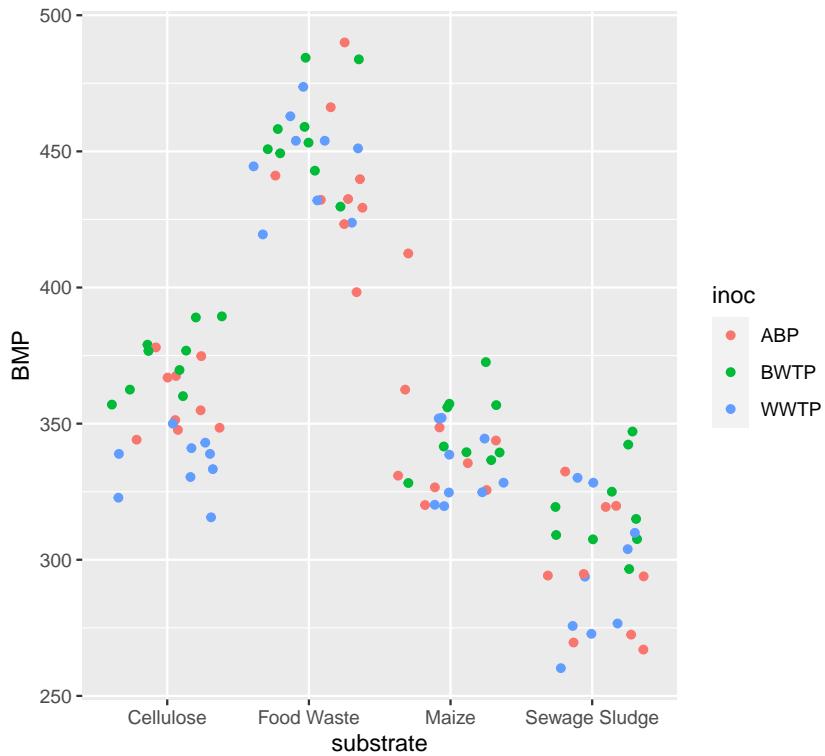
```
#      substrate inoc  BMP1  BMP2  BMP3  BMP4  BMP5  BMP6  BMP7  BMP8
# 1  Sewage Sludge WWTP 293.8 272.8 303.9 260.2 275.7 276.6 309.9 330.1
# 2          Maize WWTP 319.7 320.2 344.5 324.7 328.3 338.6 324.8 351.9
# 3      Food Waste WWTP 453.9 444.5 462.9 451.1 453.9 473.7 423.8 419.5
# 4      Cellulose WWTP 333.3 315.6 341.0 322.8 330.4 338.9 338.9 343.0
# 5  Sewage Sludge ABP 294.8 294.2 293.9 267.0 269.6 272.5 332.4 319.8
# 6          Maize ABP 320.1 325.6 348.6 362.5 343.8 412.5 326.6 330.9
# 7      Food Waste ABP 441.1 432.2 466.2 490.0 398.3 429.3 423.3 432.5
# 8      Cellulose ABP 344.1 347.7 374.8 348.5 351.3 378.0 354.9 367.5
# 9  Sewage Sludge BWTP 296.6 307.6 307.5 309.1 315.0 319.4 342.3 325.0
# 10         Maize BWTP 328.2 341.6 356.8 339.4 357.3 372.6 336.6 339.5
# 11      Food Waste BWTP 459.0 450.8 484.4 453.2 449.3 483.8 442.9 429.7
# 12      Cellulose BWTP 379.0 389.4 376.8 360.1 357.0 389.0 362.5 369.7
#      BMP9
# 1  328.3
# 2  352.1
# 3  432.0
# 4  350.0
# 5  319.4
# 6  335.5
# 7  439.8
# 8  366.9
# 9  347.1
# 10 356.0
# 11 458.2
# 12 376.7
```

This structure could work well in a spreadsheet analysis. For analysis in R, the structure can be changed to long using the `gather()` function.

```
bil <- gather(bi, key = 'rep', value = 'BMP', contains('BMP'))  
head(bil)  
  
#      substrate inoc rep BMP  
# 1 Sewage Sludge WWTP BMP1 293.8  
# 2 Maize WWTP BMP1 319.7  
# 3 Food Waste WWTP BMP1 453.9  
# 4 Cellulose WWTP BMP1 333.3  
# 5 Sewage Sludge ABP BMP1 294.8  
# 6 Maize ABP BMP1 320.1  
  
dim(bil)  
  
# [1] 108 4  
  
dfsumm(bil)  
  
#  
# 108 rows and 4 columns  
# 108 unique rows  
#  
#          substrate      inoc      rep      BMP  
# Class      character   character   character numeric  
# Minimum    Cellulose     ABP       BMP1      260  
# Maximum    Sewage Sludge   WWTP      BMP9      490  
# Mean        Food Waste    BWTP      BMP5      362  
# Unique (excl. NA)      4         3         9      103  
# Missing values      0         0         0         0  
# Sorted       FALSE      FALSE      TRUE      FALSE
```

Here are the values, with a single point representing a BMP value from a single bottle.

```
ggplot(bil, aes(substrate, BMP, colour = inoc)) +  
  geom_jitter(height = 0)
```



Calculate means and standard deviation.

```
bm <- as.data.frame(summarise(group_by(bil, substrate, inoc), BMP.mn = mean(BMP),
                                BMP.sd = sd(BMP), n = length(BMP)))

# `summarise()` has grouped output by 'substrate'. You can override using the '.groups'
# argument.

bm$BMP.se = bm$BMP.sd / sqrt(bm$n)

bm

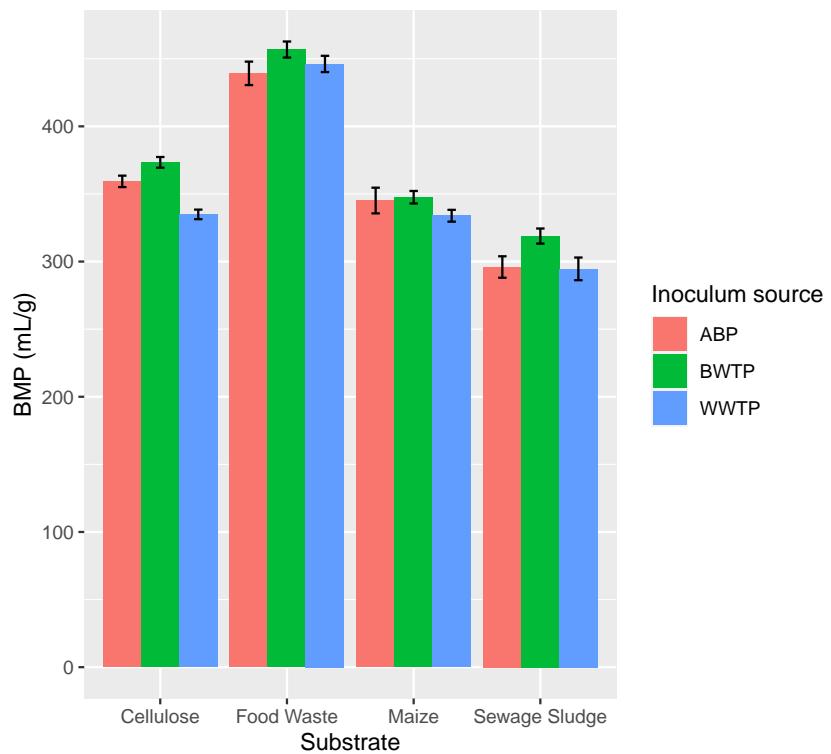
#   substrate inoc    BMP.mn    BMP.sd n    BMP.se
# 1 Cellulose ABP 359.3000 12.65178 9 4.217260
# 2 Cellulose BWTP 373.3556 11.89276 9 3.964254
# 3 Cellulose WWTP 334.8778 10.63329 9 3.544431
# 4 Food Waste ABP 439.1889 26.05554 9 8.685180
# 5 Food Waste BWTP 456.8111 17.78479 9 5.928262
# 6 Food Waste WWTP 446.1444 18.01694 9 6.005648
# 7     Maize ABP 345.1222 28.50604 9 9.502014
# 8     Maize BWTP 347.5556 13.87661 9 4.625536
# 9     Maize WWTP 333.8667 13.12355 9 4.374516
# 10 Sewage Sludge ABP 295.9556 23.81765 9 7.939215
# 11 Sewage Sludge BWTP 318.8444 16.75717 9 5.585724
# 12 Sewage Sludge WWTP 294.5889 25.14202 9 8.380673
```

And plot them.

```

ggplot(bm, aes(substrate, BMP.mn, fill = inoc)) +
  geom_bar(position = position_dodge(), stat = 'identity') +
  geom_errorbar(aes(ymin = BMP.mn - BMP.se, ymax = BMP.mn + BMP.se), position = position_dodge(0.9),
  labs(x = 'Substrate', y = 'BMP (mL/g)', fill = 'Inoculum source')

```



Here is a case where we really do need a statistical analysis to help understand the data.

```

m1 <- lm(BMP ~ substrate * inoc, data = bil)
summary(m1)

#
# Call:
# lm(formula = BMP ~ substrate * inoc, data = bil)
#
# Residuals:
#      Min       1Q   Median       3Q      Max 
# -40.889  -11.719   -1.700    9.261   67.378 
#
# Coefficients:
#                               Estimate Std. Error t value Pr(>|t|)    
# (Intercept)                  359.300    6.377  56.343 < 2e-16  
# substrateFood Waste           79.889    9.018   8.858 4.21e-14  
# substrateMaize                -14.178   9.018  -1.572  0.11922  
# substrateSewage Sludge        -63.344   9.018  -7.024 3.10e-10  
# inocBWTP                      14.056    9.018   1.559  0.12240  
# inocWWTP                      -24.422   9.018  -2.708  0.00801  
# substrateFood Waste:inocBWTP     3.567   12.754   0.280  0.78035 

```

```

# substrateMaize:inocBWTP      -11.622    12.754  -0.911  0.36444
# substrateSewage Sludge:inocBWTP   8.833    12.754   0.693  0.49024
# substrateFood Waste:inocWWTP     31.378    12.754   2.460  0.01567
# substrateMaize:inocWWTP        13.167    12.754   1.032  0.30450
# substrateSewage Sludge:inocWWTP   23.056    12.754   1.808  0.07378
#
# (Intercept)                  ***
# substrateFood Waste           ***
# substrateMaize
# substrateSewage Sludge       ***
# inocBWTP
# inocWWTP                     **
# substrateFood Waste:inocBWTP
# substrateMaize:inocBWTP
# substrateSewage Sludge:inocBWTP
# substrateFood Waste:inocWWTP   *
# substrateMaize:inocWWTP
# substrateSewage Sludge:inocWWTP .
# ---
# Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#
# Residual standard error: 19.13 on 96 degrees of freedom
# Multiple R-squared:  0.8995, Adjusted R-squared:  0.888
# F-statistic: 78.14 on 11 and 96 DF, p-value: < 2.2e-16

```

`anova(m1)`

```

# Analysis of Variance Table
#
# Response: BMP
#              Df Sum Sq Mean Sq F value    Pr(>F)
# substrate      3 302030 100677 275.0758 < 2.2e-16 ***
# inoc          2   8804    4402 12.0276 2.181e-05 ***
# substrate:inoc 6   3740     623  1.7031    0.1285
# Residuals     96  35136     366
# ---
# Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

There is clear evidence of an inoculum effect, and a slight suggestion of a possible interaction.

```

m2 <- aov(BMP ~ substrate * inoc, data = bil)
summary(m2)

#
#              Df Sum Sq Mean Sq F value    Pr(>F)
# substrate      3 302030 100677 275.076 < 2e-16 ***
# inoc          2   8804    4402 12.028 2.18e-05 ***
# substrate:inoc 6   3740     623  1.703    0.129
# Residuals     96  35136     366
# ---
# Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

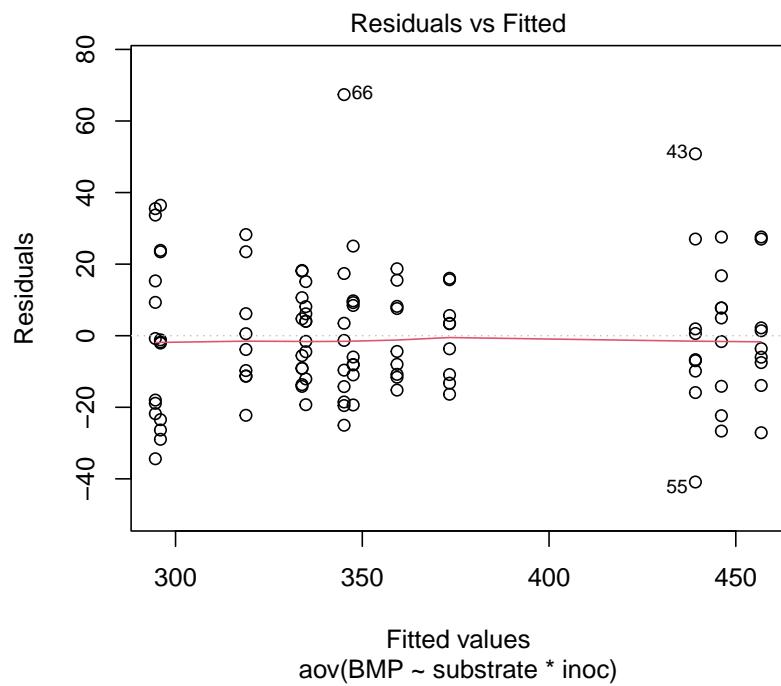
```

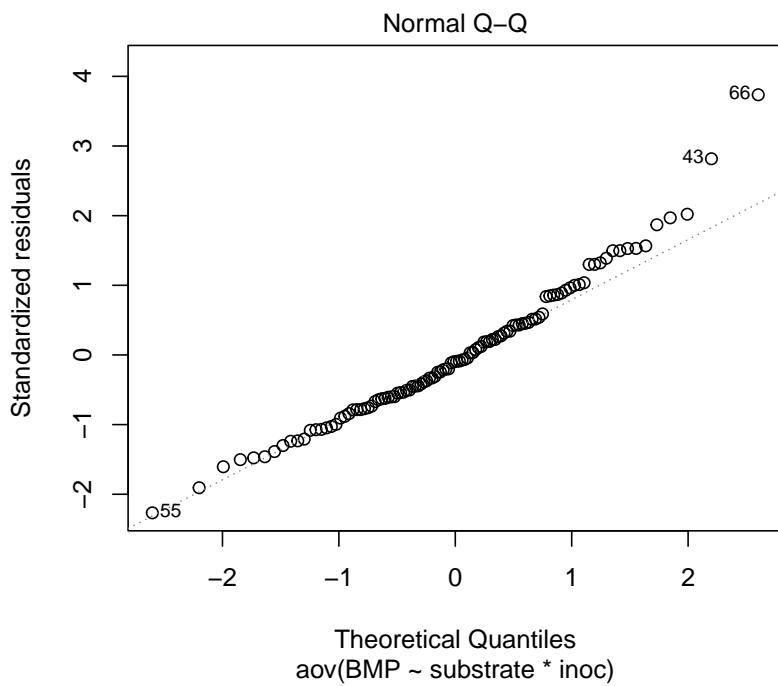
TukeyHSD(m2, 'inoc')

# Tukey multiple comparisons of means
# 95% family-wise confidence level
#
# Fit: aov(formula = BMP ~ substrate * inoc, data = bil)
#
# $inoc
#          diff      lwr      upr     p adj
# BWTP-ABP 14.250000  3.515301 24.984699 0.0059271
# WWTP-ABP -7.522222 -18.256921  3.212477 0.2227058
# WWTP-BWTP -21.772222 -32.506921 -11.037523 0.0000154

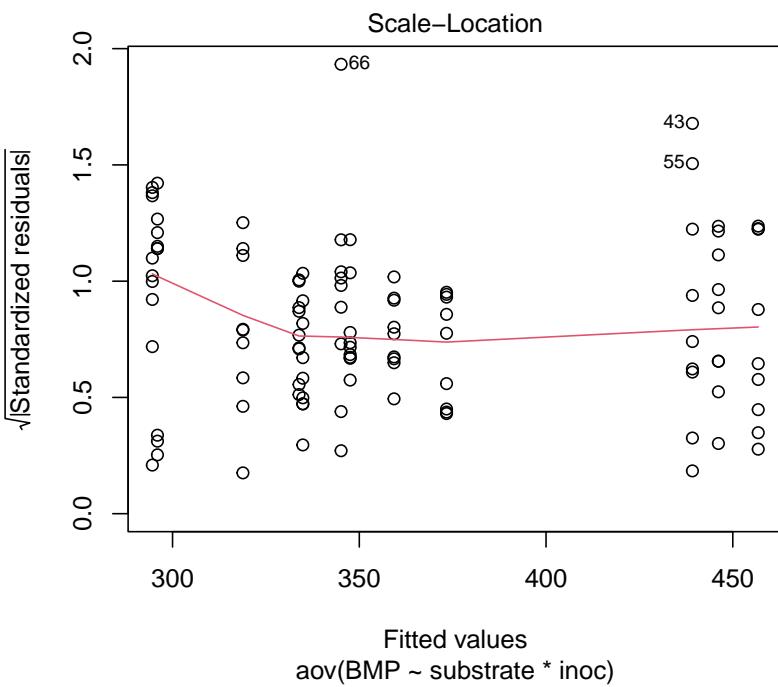
```

```
plot(m2, ask = FALSE)
```





```
# hat values (leverages) are all = 0.1111111
# and there are no factor predictors; no plot no. 5
```



```

m3 <- aov(log10(BMP) ~ substrate * inoc, data = bil)
summary(m3)

#               Df Sum Sq Mean Sq F value    Pr(>F)
# substrate      3 0.4081 0.13604 244.417 < 2e-16 ***
# inoc          2 0.0141 0.00703 12.623 1.36e-05 ***
# substrate:inoc 6 0.0062 0.00103   1.853    0.097 .
# Residuals     96 0.0534 0.00056
# ---
# Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(tr <- TukeyHSD(m3, 'inoc'))

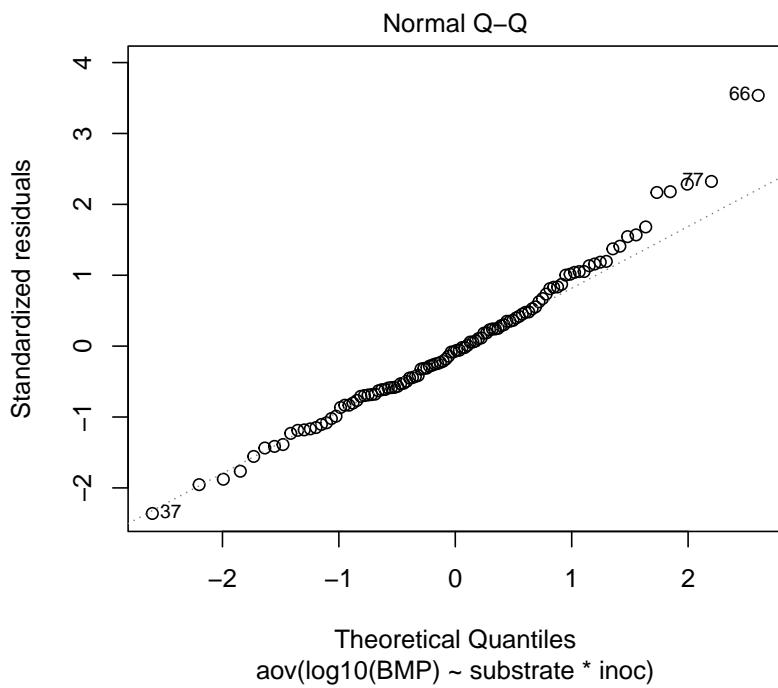
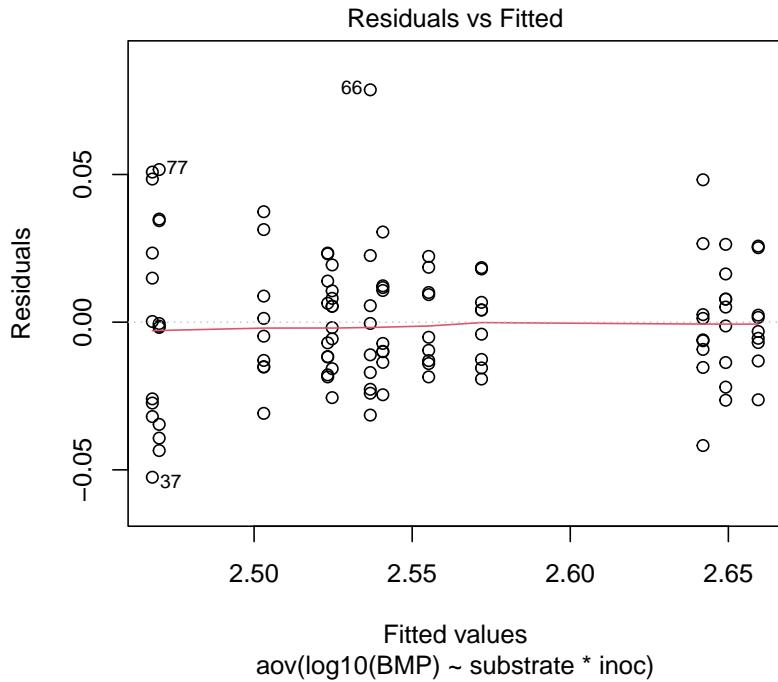
# Tukey multiple comparisons of means
# 95% family-wise confidence level
#
# Fit: aov(formula = log10(BMP) ~ substrate * inoc, data = bil)
#
# $inoc
#           diff       lwr       upr     p adj
# BWTP-ABP  0.017803269  0.004565351  0.03104119 0.0052233
# WWTP-ABP -0.009747578 -0.022985495  0.00349034 0.1911260
# WWTP-BWTP -0.027550847 -0.040788764 -0.01431293 0.0000092

100 * (10^tr$inoc[, 'diff'] - 1)

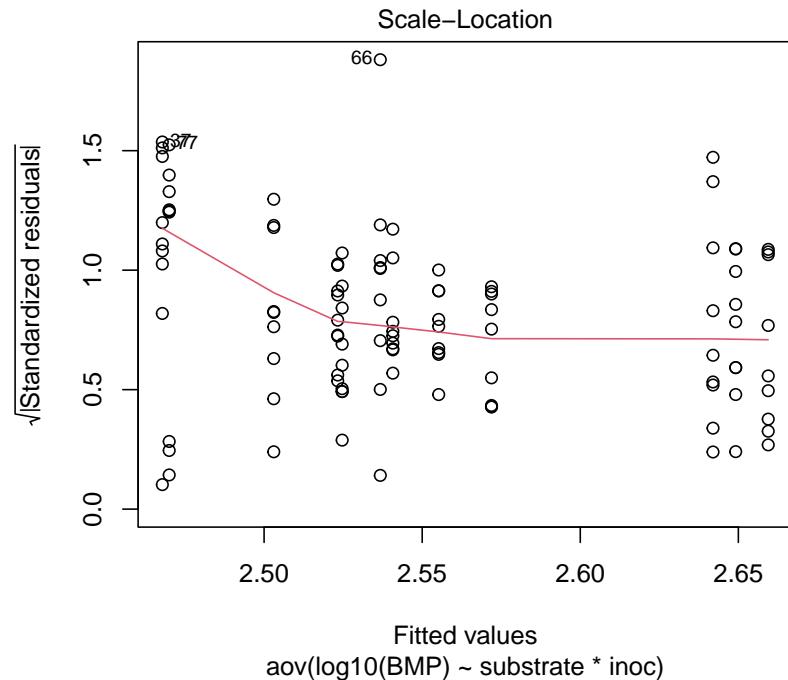
#  BWTP-ABP  WWTP-ABP  WWTP-BWTP
# 4.184538 -2.219462 -6.146785

```

```
plot(m3, ask = FALSE)
```



```
# hat values (leverages) are all = 0.1111111  
# and there are no factor predictors; no plot no. 5
```



We can conclude that the BWTP inoculum resulted in BMP values about 4-6% higher than the other two.

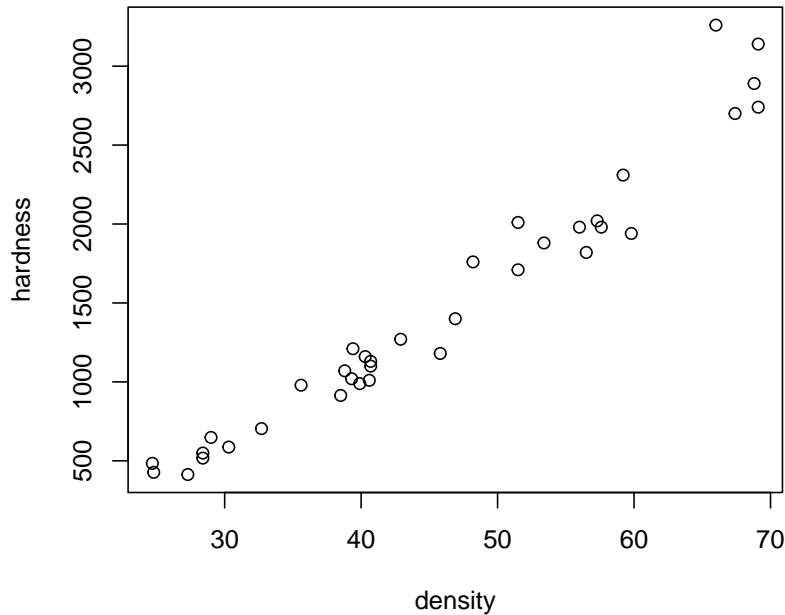
3 Problem 2. Wood hardness and density

```
hard <- read.csv("data/janka.csv")
dfsumm(hard)

#
# 36 rows and 2 columns
# 36 unique rows
#          density hardness
# Class      numeric   integer
# Minimum     24.7      413
# Maximum     69.1     3260
# Mean        45.7     1180
# Unique (excl. NA) 32       35
# Missing values 0       0
# Sorted      TRUE    FALSE
```

Let's start out by seeing what the data look like.

```
plot(hardness ~ density, data = hard)
```



We might be interested in doing two things with these data: determining if wood hardness (difficult to measure) is related to wood density (easy to measure), and, if so, predicting hardness from the

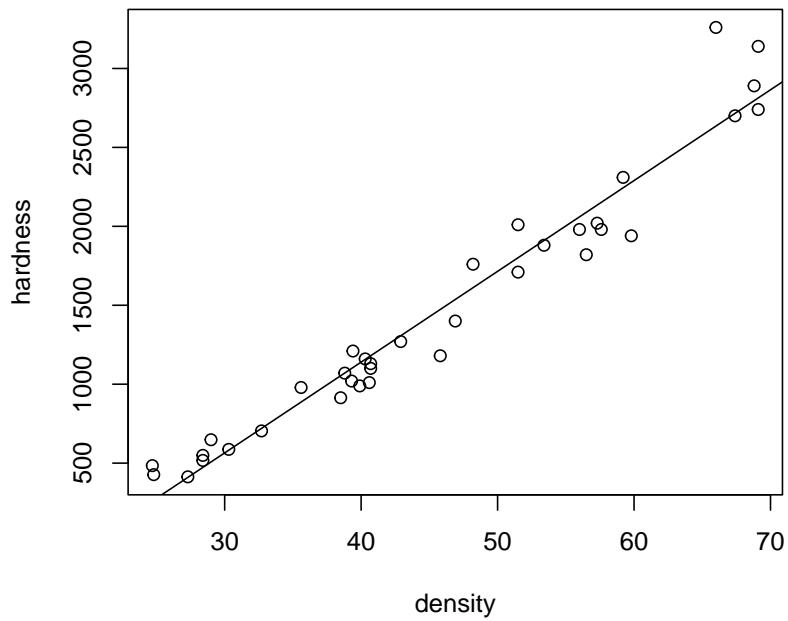
density. Are these data experimental or observational? Try to fit an appropriate regression model to these data, and take a look at the residuals to check the structure. Can you improve it?

```
m1 <- lm(hardness ~ density, data = hard)
summary(m1)

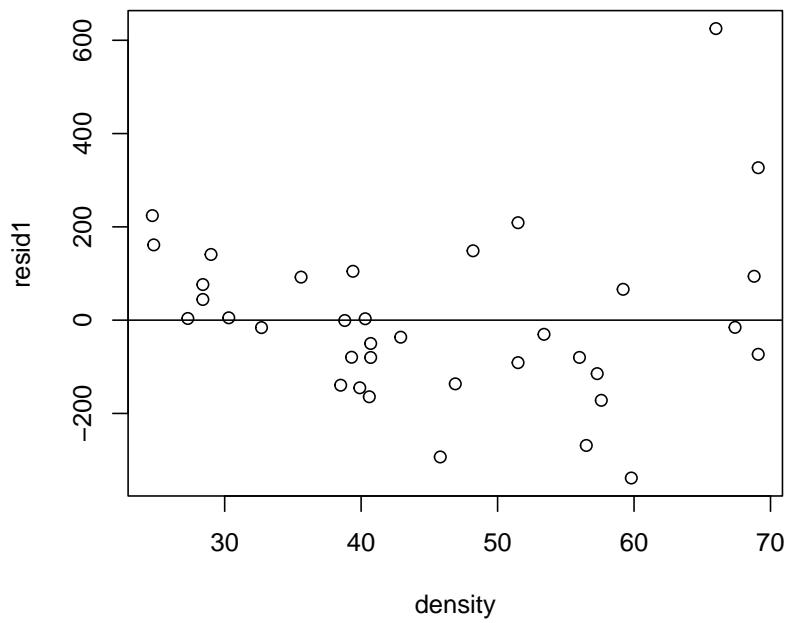
#
# Call:
# lm(formula = hardness ~ density, data = hard)
#
# Residuals:
#     Min      1Q  Median      3Q     Max 
# -338.40  -96.98  -15.71   92.71  625.06 
#
# Coefficients:
#             Estimate Std. Error t value Pr(>|t|)    
# (Intercept) -1160.500    108.580  -10.69 2.07e-12 ***
# density       57.507      2.279   25.24 < 2e-16 ***
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#
# Residual standard error: 183.1 on 34 degrees of freedom
# Multiple R-squared:  0.9493, Adjusted R-squared:  0.9478 
# F-statistic:  637 on 1 and 34 DF,  p-value: < 2.2e-16

hard$pred1 <- predict(m1)
hard$resid1 <- resid(m1)

plot(hardness ~ density, data = hard)
abline(m1)
```



```
plot(resid1 ~ density, data = hard)
abline(h = 0)
```



```

m2 <- lm(hardness ~ poly(density, 3), data = hard)
summary(m2)

#
# Call:
# lm(formula = hardness ~ poly(density, 3), data = hard)
#
# Residuals:
#      Min       1Q   Median       3Q      Max
# -310.98  -92.52  -14.94   61.41  537.92
#
# Coefficients:
#              Estimate Std. Error t value Pr(>|t|)
# (Intercept) 1469.47    27.29  53.841 < 2e-16 ***
# poly(density, 3)1 4620.14   163.76  28.213 < 2e-16 ***
# poly(density, 3)2  525.40   163.76   3.208  0.00303 **
# poly(density, 3)3   72.14   163.76   0.441  0.66252
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

#
# Residual standard error: 163.8 on 32 degrees of freedom
# Multiple R-squared:  0.9618, Adjusted R-squared:  0.9583
# F-statistic: 268.8 on 3 and 32 DF,  p-value: < 2.2e-16

```

```

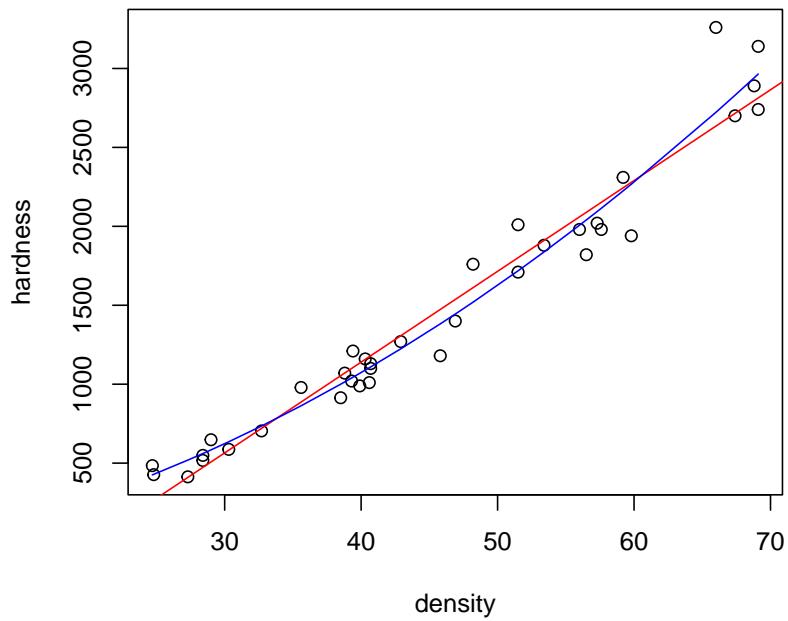
m2 <- lm(hardness ~ poly(density, 2), data = hard)
hard$pred2 <- predict(m2)
hard$resid2 <- resid(m2)

```

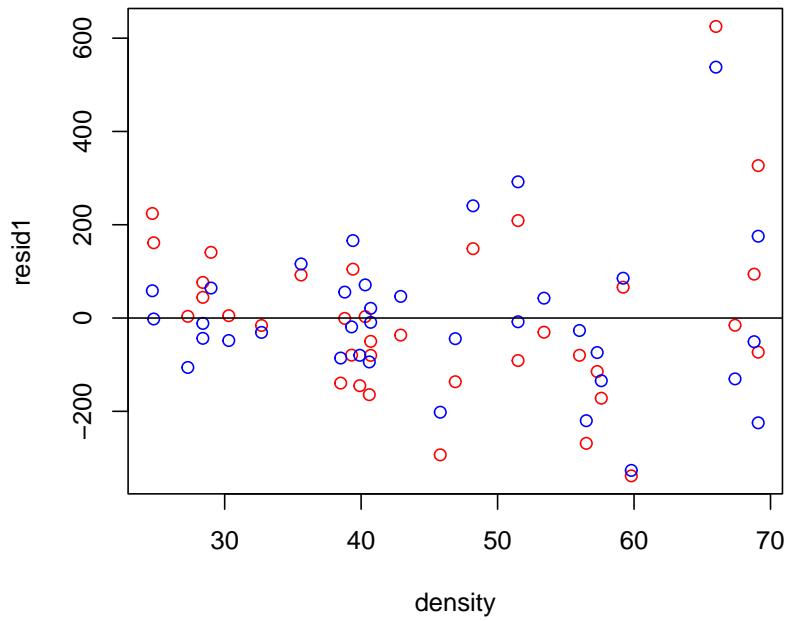
```

plot(hardness ~ density, data = hard)
abline(m1, col = 'red')
lines(pred2 ~ density, data = hard, col = 'blue')

```



```
plot(resid1 ~ density, data = hard, col = 'red')
points(resid2 ~ density, data = hard, col = 'blue')
abline(h = 0)
```



```

m3 <- lm(log10(hardness) ~ poly(density, 2), data = hard)
summary(m3)

#
# Call:
# lm(formula = log10(hardness) ~ poly(density, 2), data = hard)
#
# Residuals:
#       Min     1Q Median     3Q    Max
# -0.096983 -0.024792 -0.004795  0.032573  0.081955
#
# Coefficients:
#                   Estimate Std. Error t value Pr(>|t|)
# (Intercept)      3.099195   0.007294 424.896 < 2e-16 ***
# poly(density, 2)1 1.470617   0.043764  33.603 < 2e-16 ***
# poly(density, 2)2 -0.234322   0.043764  -5.354 6.49e-06 ***
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

#
# Residual standard error: 0.04376 on 33 degrees of freedom
# Multiple R-squared:  0.9723, Adjusted R-squared:  0.9706
# F-statistic: 578.9 on 2 and 33 DF,  p-value: < 2.2e-16

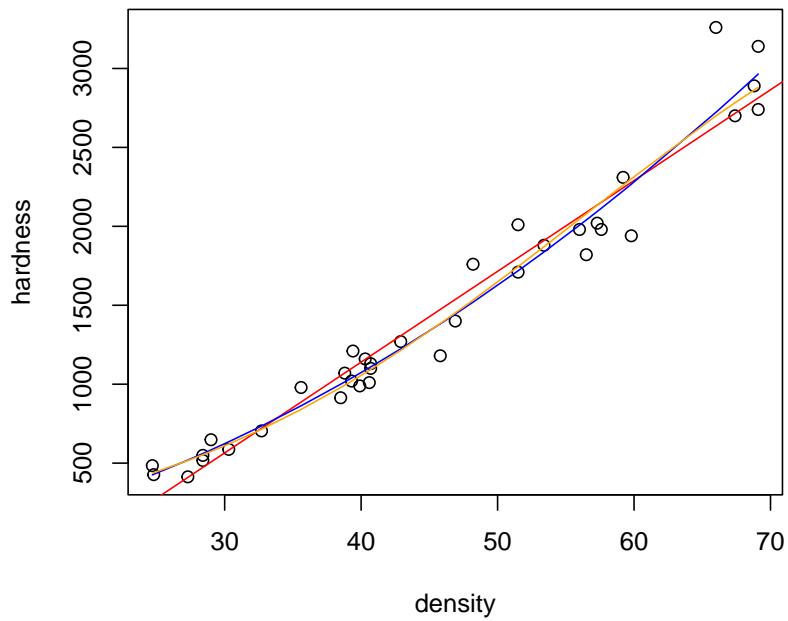
hard$pred3 <- 10^predict(m3)
hard$resid3 <- hard$pred3 - hard$hardness

```

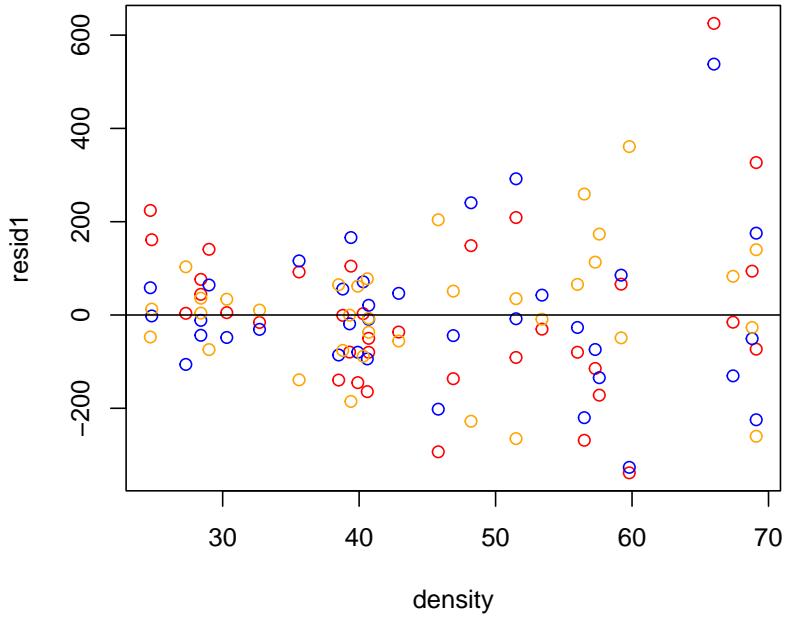
```

plot(hardness ~ density, data = hard)
abline(m1, col = 'red')
lines(pred2 ~ density, data = hard, col = 'blue')
lines(pred3 ~ density, data = hard, col = 'orange')

```



```
plot(resid1 ~ density, data = hard, col = 'red')
points(resid2 ~ density, data = hard, col = 'blue')
points(resid3 ~ density, data = hard, col = 'orange')
abline(h = 0)
```



4 Problem 3. Fruit fly longevity and sexual activity

The data in the file fruitfly.csv are from an experiment on fruitfly longevity and are also from Faraway [2005]. The original objective of this famous experiment was to assess the effect of sexual activity (manipulated by controlling the number of females placed with a single male, `activity` column) on fruitfly longevity (how long the flies live, `longevity` column). But longevity is known to be correlated with thorax length (`thorax` column).

```
ff <- read.csv('data/fruitfly.csv')
head(ff)

#   thorax longevity activity
# 1  0.68       37    many
# 2  0.68       49    many
# 3  0.72       46    many
# 4  0.72       63    many
# 5  0.76       39    many
# 6  0.76       46    many
```

1. How might you plot these data to assess the effect of activity?
2. How can you fit a statistical model that utilizes the correlation with thorax length to increase power?
3. What approach should you use to compare the levels of `activity` to each other?

```

ggplot(ff, aes(thorax, longevity, colour = activity)) +
  geom_point() +
  geom_smooth(se = FALSE) +
  labs(x = 'Thorax length (mm)', y = 'Longevity (days)', colour = 'Activity')

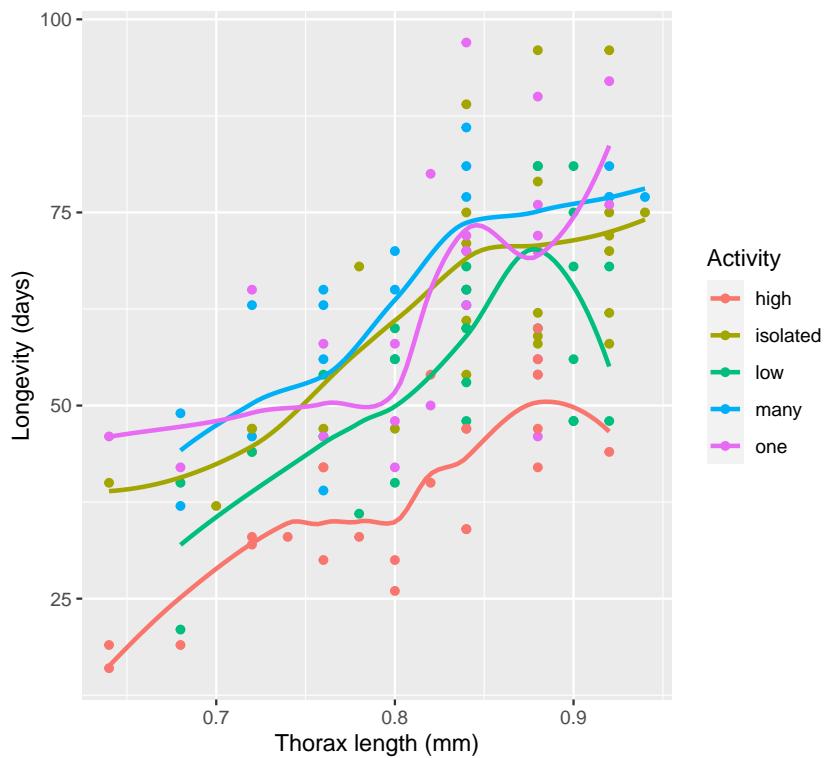
# 'geom_smooth()' using method = 'loess' and formula 'y ~ x'

# Warning in simpleLoess(y, x, w, span, degree = degree, parametric = parametric, :
# pseudoinverse used at 0.84

# Warning in simpleLoess(y, x, w, span, degree = degree, parametric = parametric, :
# neighborhood radius 0.04

# Warning in simpleLoess(y, x, w, span, degree = degree, parametric = parametric, :
# reciprocal condition number  8.6863e-22

```

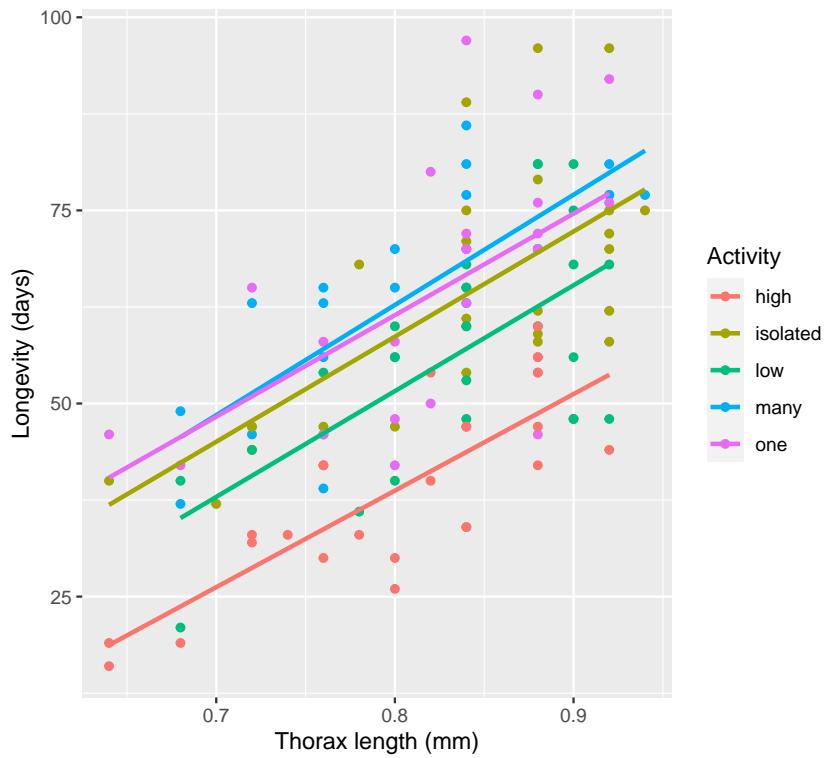


```

ggplot(ff, aes(thorax, longevity, colour = activity)) +
  geom_point() +
  geom_smooth(method = lm, se = FALSE) +
  labs(x = 'Thorax length (mm)', y = 'Longevity (days)', colour = 'Activity')

# 'geom_smooth()' using formula 'y ~ x'

```

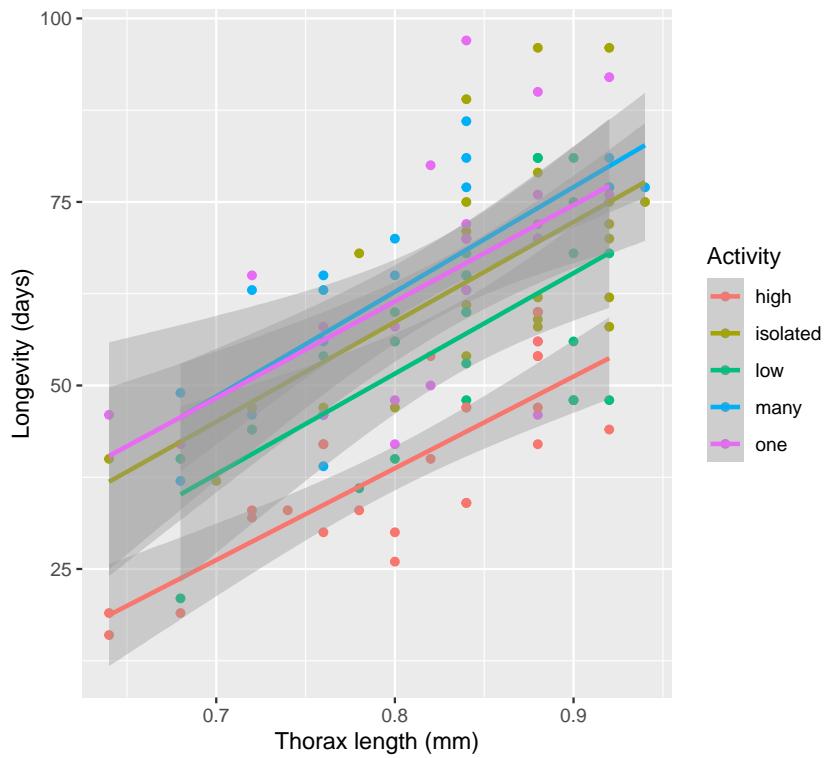


```

ggplot(ff, aes(thorax, longevity, colour = activity)) +
  geom_point() +
  geom_smooth(method = lm) +
  labs(x = 'Thorax length (mm)', y = 'Longevity (days)', colour = 'Activity')

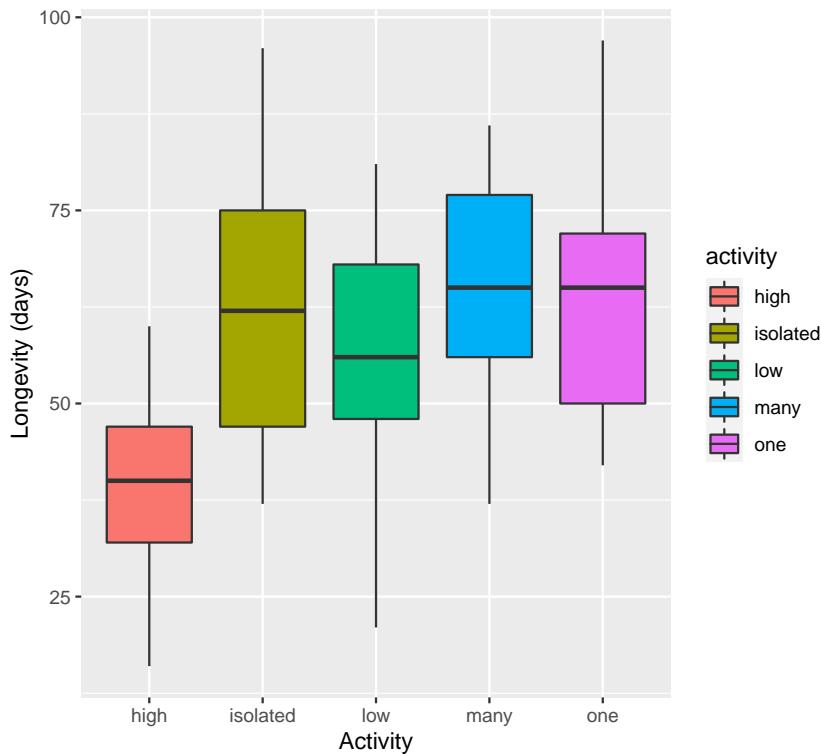
# 'geom_smooth()' using formula 'y ~ x'

```



Compare to a boxplot—more difficult to see separation of groups.

```
ggplot(ff, aes(activity, longevity, fill = activity)) +
  geom_boxplot() +
  labs(x = 'Activity', y = 'Longevity (days)', colour = 'Activity')
```



```
levels(ff$activity)
```

```
# NULL
```

First level will be reference. Let's change it to isolated.

```
ff$activity <- relevel(ff$activity, ref= 'isolated')
```

```
# Error in relevel.default(ff$activity, ref = "isolated"): 'relevel' only for (unordered)
# factors
```

```
m1 <- lm(longevity ~ activity * thorax, data = ff)
anova(m1)
```

```
# Analysis of Variance Table
#
# Response: longevity
#              Df  Sum Sq Mean Sq F value    Pr(>F)
# activity       4 12269.5  3067.4  26.728 1.2e-15 ***
# thorax         1 12368.4 12368.4 107.774 < 2e-16 ***
# activity:thorax 4    24.3     6.1   0.053  0.9947
# Residuals     114 13083.0   114.8
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```

m2 <- lm(longevity ~ activity + thorax, data = ff)
anova(m2)

# Analysis of Variance Table
#
# Response: longevity
#           Df Sum Sq Mean Sq F value    Pr(>F)
# activity      4   12270  3067.4  27.614 3.481e-16 ***
# thorax        1   12368 12368.4 111.348 < 2.2e-16 ***
# Residuals  118   13107   111.1
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

summary(m2)

#
# Call:
# lm(formula = longevity ~ activity + thorax, data = ff)
#
# Residuals:
#       Min     1Q Median     3Q    Max
# -26.108 -7.014 -1.101  6.234 30.265
#
# Coefficients:
#             Estimate Std. Error t value Pr(>|t|)    
# (Intercept) -68.753     10.401  -6.610 1.17e-09 ***
# activityisolated 20.004      3.016   6.632 1.05e-09 ***
# activitylow     12.989      3.019   4.302 3.51e-05 ***
# activitymany    24.142      3.016   8.005 9.38e-13 ***
# activityone     22.641      2.999   7.550 1.01e-11 ***
# thorax         134.341     12.731  10.552 < 2e-16 ***
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#
# Residual standard error: 10.54 on 118 degrees of freedom
# Multiple R-squared:  0.6527, Adjusted R-squared:  0.638 
# F-statistic: 44.36 on 5 and 118 DF,  p-value: < 2.2e-16

```

We can use Bonferroni adjustment, $0.05 / 5 = 0.01$. So only `high` level is clearly different—20 days shorter longevity, which is a lot!

```

confint(m2)

#
#           2.5 %    97.5 %
# (Intercept) -89.349458 -48.15674
# activityisolated 14.031174 25.97625
# activitylow     7.009965 18.96756
# activitymany    18.169733 30.11507
# activityone     16.702511 28.57921
# thorax         109.130197 159.55255

```

Strange that “many” level is so different from others.

```
confint(m2)
```

```
#               2.5 %    97.5 %
# (Intercept) -89.349458 -48.15674
# activityisolated 14.031174 25.97625
# activitylow      7.009965 18.96756
# activitymany     18.169733 30.11507
# activityone      16.702511 28.57921
# thorax          109.130197 159.55255
```

5 Problem 4: Growth and nitrate accumulation by *Lemna minor*

Duckweeds are very tiny floating plants that can be used for wastewater treatment and recovery of nitrogen. Harvested material can be used as an animal feed. Devlamynck et al. [2020] measured biomass production and nitrate accumulation in a duckweed species *Lemna minor*. The data are in lemma.csv. Use them to explore the following questions.

1. Did medium affect growth (`grow`)?
2. Did medium affect NO_3^- accumulation (`NO3.accum`)?
3. Is NO_3^- accumulation related to NO_3^- concentration in the medium (`NO3.med`)?

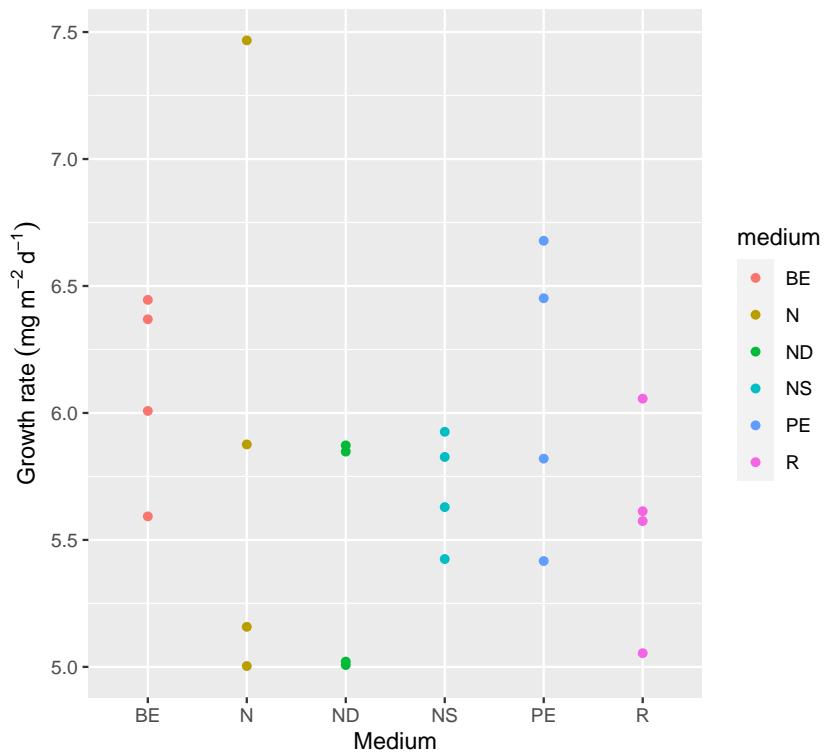
```
lem <- read.csv('data/lemma.csv')
```

```
summary(lem)
```

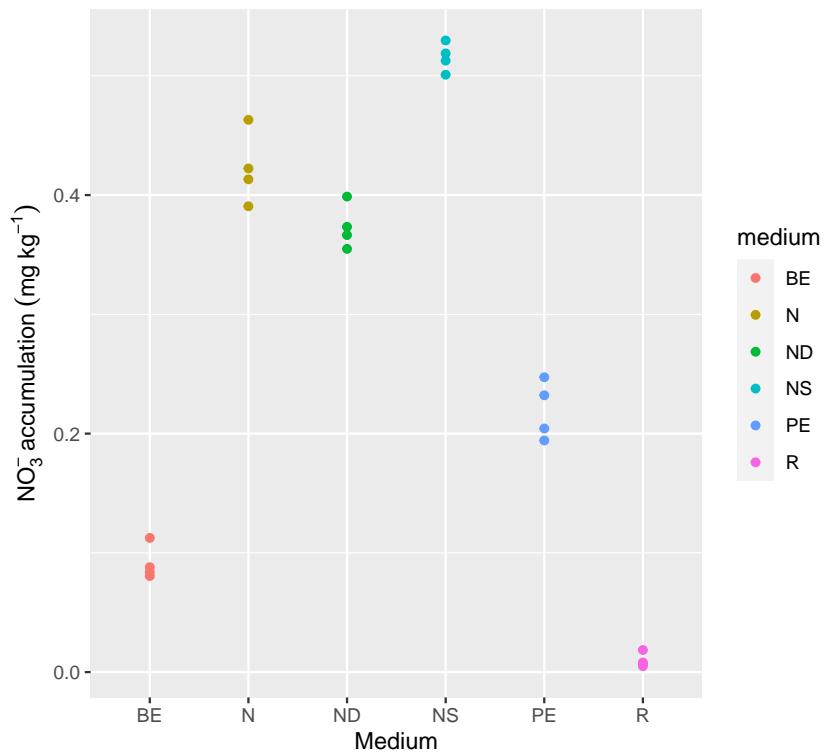
```
#   med.descrip           medium            grow
#   Length:24           Length:24       Min.   :5.003
#   Class  :character   Class  :character  1st Qu.:5.423
#   Mode   :character   Mode   :character  Median :5.823
#                                         Mean   :5.797
#                                         3rd Qu.:6.020
#                                         Max.   :7.467
#   NO3.accum           pH.med        NO3.med
#   Min.   :0.005025   Min.   :5.760   Min.   : 0.009594
#   1st Qu.:0.087042   1st Qu.:6.388   1st Qu.: 2.215323
#   Median :0.301076   Median :7.390   Median : 4.138554
#   Mean   :0.271930   Mean   :7.461   Mean   : 7.795348
#   3rd Qu.:0.415473   3rd Qu.:8.525   3rd Qu.: 9.410879
#   Max.   :0.529639   Max.   :9.632   Max.   :27.129694
```

```
library(ggplot2)
```

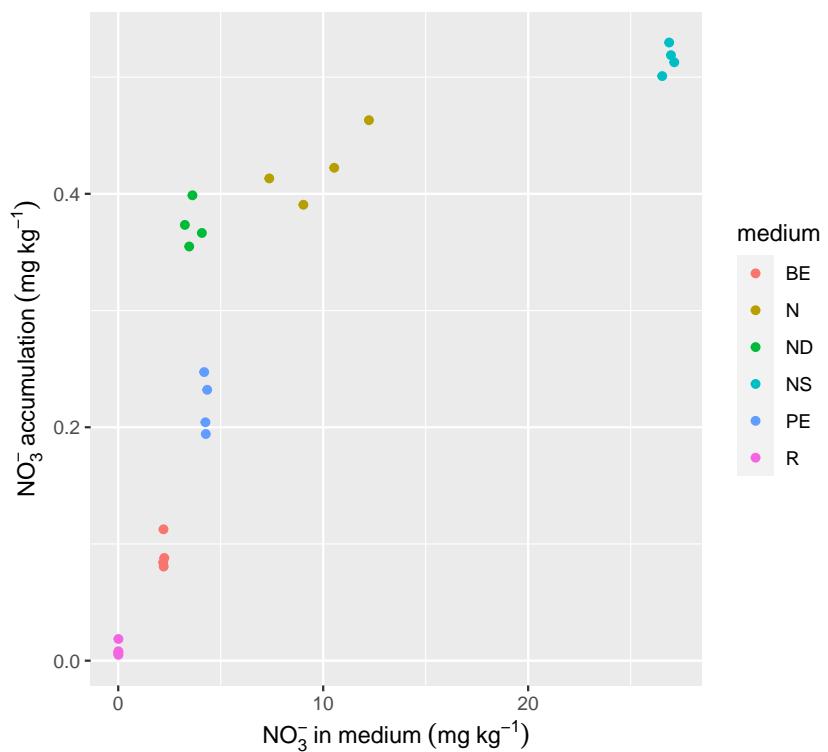
```
ggplot(lem, aes(medium, grow, colour = medium)) +  
  geom_point() +  
  labs(x = 'Medium', y = expression('Growth rate'^(mg~m^-2~d^-1)))
```



```
ggplot(lem, aes(medium, NO3.accum, colour = medium)) +  
  geom_point() +  
  labs(x = 'Medium', y = expression(NO[3]^-^- accumulation^(mg~kg^-1)))
```



```
ggplot(lem, aes(NO3.med, NO3.accum, colour = medium)) +
  geom_point() +
  labs(x = expression(NO[3]^'-'~'in medium'^(mg~kg^-1)), y = expression(NO[3]^'-'~'accumulation'^(mg~kg^-1)))
```



First growth. Check plot—no clear effect, no stats needed. We can calculate average and sd at least.

```
lemsum <- as.data.frame(summarise(group_by(lem, medium),
                                     grow.mean = mean(grow), grow.sd = sd(grow),
                                     N03.med.mean = mean(N03.med), N03.med.sd = sd(N03.med),
                                     N03.accum.mean = mean(N03.accum),
                                     N03.accum.sd = sd(N03.accum)))
```

```
lemsum
```

	medium	grow.mean	grow.sd	N03.med.mean	N03.med.sd	N03.accum.mean
# 1	BE	6.103780	0.3904054	2.217319070	0.0227861966	0.091324595
# 2	N	5.876119	1.1268901	9.794380141	2.0779095490	0.422308931
# 3	ND	5.437048	0.4885779	3.604850206	0.3524655526	0.373366822
# 4	NS	5.701615	0.2220019	26.879508387	0.2478468075	0.515520870
# 5	PE	6.091729	0.5780991	4.266298899	0.0601486883	0.219471297
# 6	R	5.574388	0.4101557	0.009733427	0.0001607362	0.009586386
				N03.accum.sd		
# 1		0.014458101				
# 2		0.030275823				
# 3		0.018557293				
# 4		0.011995311				
# 5		0.024499245				
# 6		0.006123297				

For nitrate accumulation, there seem to be effects.

```
levels(lem$medium)
```

```
# NULL
```

```
lem$medium <- relevel(lem$medium, ref= 'R')
```

```
# Error in relevel.default(lem$medium, ref = "R"): 'relevel' only for (unordered) factors
```

```
m1 <- lm(N03.accum ~ medium, data = lem)
anova(m1)
```

```
# Analysis of Variance Table
#
# Response: N03.accum
#           Df  Sum Sq  Mean Sq F value    Pr(>F)
# medium      5 0.78574 0.157147  418.76 < 2.2e-16 ***
# Residuals 18 0.00675 0.000375
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(m1)
```

```

#
# Call:
# lm(formula = N03.accum ~ medium, data = lem)
#
# Residuals:
#       Min     1Q Median     3Q    Max
# -0.031673 -0.009509 -0.002861  0.009905  0.040788
#
# Coefficients:
#             Estimate Std. Error t value Pr(>|t|)
# (Intercept) 0.091325  0.009686  9.429 2.19e-08 ***
# mediumN     0.330984  0.013698 24.163 3.60e-15 ***
# mediumND    0.282042  0.013698 20.590 5.83e-14 ***
# mediumNS    0.424196  0.013698 30.968 < 2e-16 ***
# mediumPE    0.128147  0.013698  9.355 2.47e-08 ***
# mediumR     -0.081738  0.013698 -5.967 1.21e-05 ***
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
#
# Residual standard error: 0.01937 on 18 degrees of freedom
# Multiple R-squared:  0.9915, Adjusted R-squared:  0.9891
# F-statistic: 418.8 on 5 and 18 DF,  p-value: < 2.2e-16

```

As expected, very clear differences. Does it matter exactly which ones differed? Seems everything was higher than R.

```

m2 <- aov(N03.accum ~ medium, data = lem)
anova(m2)

# Analysis of Variance Table
#
# Response: N03.accum
#           Df  Sum Sq  Mean Sq F value    Pr(>F)
# medium      5 0.78574 0.157147 418.76 < 2.2e-16 ***
# Residuals 18 0.00675 0.000375
# ---
# Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

TukeyHSD(m2)

# Tukey multiple comparisons of means
# 95% family-wise confidence level
#
# Fit: aov(formula = N03.accum ~ medium, data = lem)
#
# $medium
#          diff      lwr      upr      p adj
# N-BE    0.33098434  0.28745154  0.374517136 0.00000000
# ND-BE   0.28204223  0.23850943  0.325575027 0.00000000
# NS-BE   0.42419628  0.38066347  0.467729076 0.00000000
# PE-BE   0.12814670  0.08461390  0.171679503 0.00000003

```

```

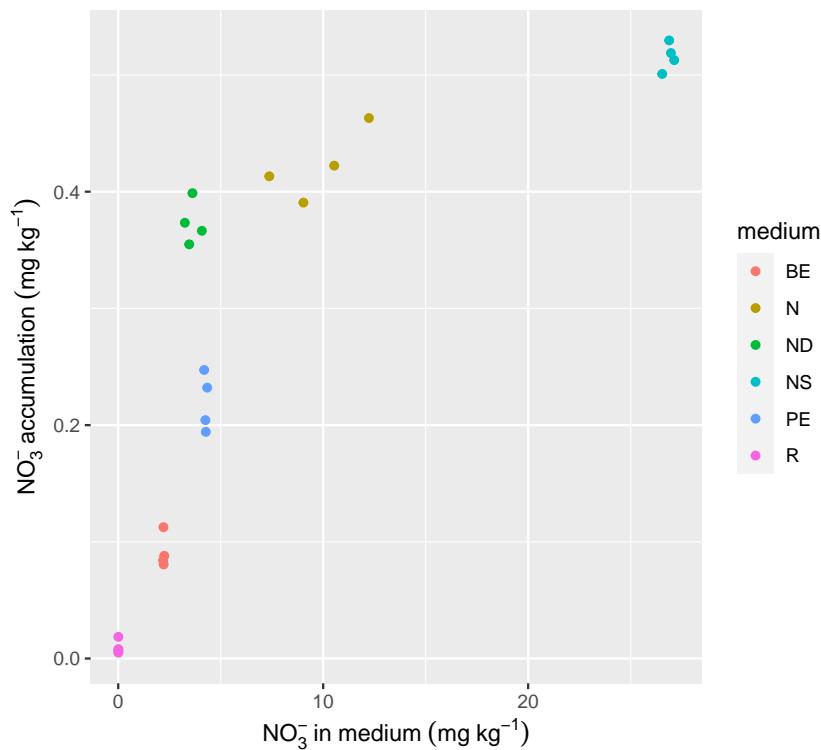
# R-BE -0.08173821 -0.12527101 -0.038205408 0.0001514
# ND-N -0.04894211 -0.09247491 -0.005409309 0.0224993
# NS-N 0.09321194 0.04967914 0.136744740 0.0000291
# PE-N -0.20283763 -0.24637043 -0.159304833 0.0000000
# R-N -0.41272254 -0.45625534 -0.369189744 0.0000000
# NS-ND 0.14215405 0.09862125 0.185686849 0.0000001
# PE-ND -0.15389552 -0.19742832 -0.110362724 0.0000000
# R-ND -0.36378044 -0.40731324 -0.320247635 0.0000000
# PE-NS -0.29604957 -0.33958237 -0.252516773 0.0000000
# R-NS -0.50593448 -0.54946728 -0.462401683 0.0000000
# R-PE -0.20988491 -0.25341771 -0.166352110 0.0000000

```

```

ggplot(lem, aes(NO3.med, NO3.accum, colour = medium)) +
  geom_point() +
  labs(x = expression(NO[3]^(-1)^~'in medium'^(mg~kg^-1)), y = expression(NO[3]^(-1)^~'accumulation'^(mg~kg^-1)))

```

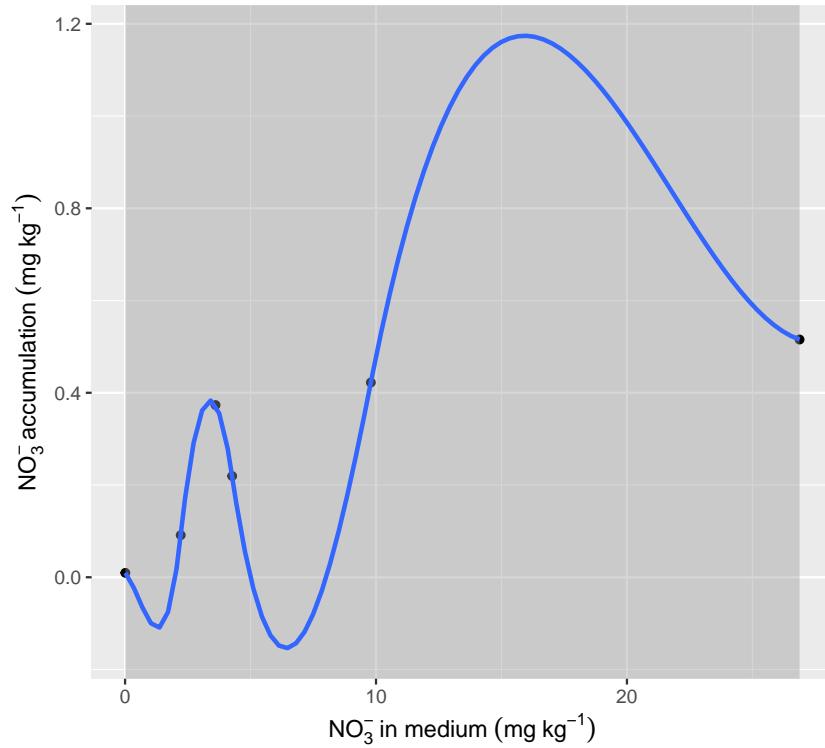


```

ggplot(lemsum, aes(NO3.med.mean, NO3.accum.mean)) +
  geom_point() +
  geom_smooth() +
  labs(x = expression(NO[3]^(-1)^~'in medium'^(mg~kg^-1)), y = expression(NO[3]^(-1)^~'accumulation'^(mg~kg^-1)))

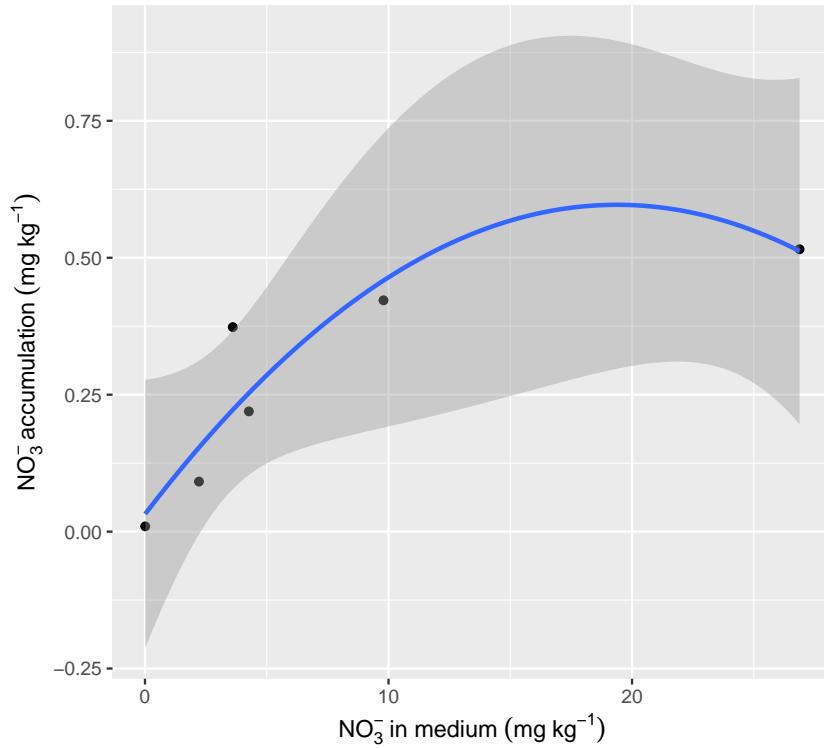
# 'geom_smooth()' using method = 'loess' and formula 'y ~ x'

```



Whoa! Overfitting by default!

```
ggplot(lmemsum, aes(NO3.med.mean, NO3.accum.mean)) +
  geom_point() +
  geom_smooth(method = lm, formula = y ~ poly(x, 2)) +
  labs(x = expression(NO[3]^'-'~'in medium'^(mg~kg^-1)), y = expression(NO[3]^'-'~'accumulation'^(mg~kg^-1)))
```



6 Bibliography

- R. Devlamynck, M. Fernandes de Souza, M. Bog, J. Leenknegt, M. Eeckhout, and E. Meers. Effect of the growth medium composition on nitrate accumulation in the novel protein crop *Lemna minor*. *Ecotoxicology and Environmental Safety*, 206:111380, Dec. 2020. ISSN 0147-6513. doi: 10.1016/j.ecoenv.2020.111380. URL <https://www.sciencedirect.com/science/article/pii/S0147651320312173>.
- J. J. Faraway. *Linear Models with R*. Number v. 63 in Texts in Statistical Science. Chapman & Hall/CRC, Boca Raton, 2005. ISBN 1-58488-425-8.
- K. Koch, T. Lippert, and J. E. Drewes. The role of inoculum's origin on the methane yield of different substrates in biochemical methane potential (BMP) tests. *Bioresource Technology*, 243 (Supplement C):457–463, Nov. 2017. ISSN 0960-8524. doi: 10.1016/j.biortech.2017.06.142.